# Trade and Industrial Policy in Supply Chains: Directed Technological Change in Rare Earths

Harald Fadinger<sup>‡</sup> Laura Alfaro\*

Jan Schymik<sup>¶</sup> Gede Virananda<sup>§</sup>

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#### Abstract

Trade and industrial policies, while primarily intended to support domestic industries, may unintentionally stimulate technological progress abroad. We document this mechanism in the case of rare earth elements (REEs) – critical inputs for manufacturing at the knowledge frontier, with low elasticity of substitution, inelastic supply, and high production and processing concentration. To assess the importance of REEs across industries, we construct an input-output table that includes disaggregated REE inputs. Using REE-related patents categorized by a large language model, sectoral TFP data, trade data, and physical and chemical substitution properties of REEs, we show that the introduction of REE export restrictions by China led to a global surge in innovation and exports in REE-intensive downstream sectors outside of China. To rationalize these findings and quantify the global impact of the adverse REE supply shock, we develop a quantitative general equilibrium model of trade and directed technological change. We also propose a structural method to estimate sectoral input substitution elasticities for REEs from patent data and find REEs to be complementary inputs. Under endogenous technologies and with complementary inputs, input supply restrictions on REEs induce a surge in REE-enhancing innovation and lead to an expansion of REE-intensive downstream sectors.

Keywords: Trade Restrictions, Industrial Policy, Global Value Chains, Rare Earths, Directed

Technological Change, Input-Output Linkages, Downstream Sectors, Innovation

JEL Codes: F13, F14, F42, O33, O47

\*Harvard Business School, NBER and CEPR. Email: lalfaro@hbs.edu; <sup>‡</sup>University of Vienna and CEPR. Email: harald.fadinger@univie.ac.at; ¶University of Mannheim, Email: jan.schymik@uni-mannheim.de; §NYU Stern, Email: iv2111@stern.nyu.edu. We thank Costas Arkolakis, Laura Bottazi, Lorenzo Caliendo, Maggie Chen, Paola Conconi, Banu Demir, Elhanan Helpman, Ben Faber, Oleg Itskhoki, Daniel Lederman, Andrés Rodríguez-Clare, Catherine Thomas, Felix Tintelnot and seminar participants at Berkeley, Bonn-Mannheim CRC TR224, 3rd CEPII IFM Workshop, CREI@30, IfW-CEPR Conference on Geoeconomics Berlin, Oxford Trade Jamboree, PSE Workshop 'Industrial Policy in the Global Economy', ISOT Conference, 'WE\_ARE\_IN' Macroeconomics and Finance conference, World Bank Workshop on Industrial Policy, ECB, TU Vienna, CEU, Bonn, Paris School of Economics, IfW Kiel, Banque de France, KU Leuven, Yale Cowles Summer Conference. Support by the Deutsche Forschungsgemeinschaft (DFG) through CRC TR 224 (Project B07) is gratefully acknowledged. We are grateful to Florian Frühhaber, Younghun Lee, Francisco Scalco, and Athibi Sharma for their excellent research assistance.

## 1 Introduction

Governments traditionally use industrial policies to selectively promote economic activity in specific sectors. While industrial policy can aim at correcting market imperfections that misalign private and social benefits, its final effects are complex when economic sectors are interconnected through global value chains. Understanding the impact of industrial policy in this environment is particularly relevant for sectors that provide essential inputs for other sectors. In the context of geopolitical tensions and the green transition, rare earth elements (REEs) have gained increased attention. Their distinct role as critical inputs for many manufacturing goods, with few possibilities of substitution, inelastic supply, and concentration of their supply in China, has sparked a broader debate about the fragility of global supply chains.

In this paper, we provide an empirical and quantitative analysis of the effects of China's policy that restricted the global supply of REEs. We advance both reduced-form causal evidence and a general-equilibrium analysis based on a novel quantitative trade model with directed technological change and input-output (IO) linkages. Our key finding is that policies that create an adverse supply shock of essential inputs can trigger, across borders, innovation, long-term productivity growth, and reallocation of economic activity towards downstream sectors that intensively use these essential inputs. Intuitively, when inputs are gross complements, a surge in the price of an input endogenously creates technological change that is directed at increasing the efficiency of input use, potentially leading to an expansion of downstream sectors that intensively use the input at the expense of other sectors. We show that REEs, because of their particular characteristics, are susceptible to these conditions.

REEs, crucial inputs in many manufacturing products due to their chemical properties, capture our interest due to at least four characteristics. First, REEs have broad and diverse applications at the knowledge frontier across various industries, including electronics, lighting, aerospace, defense, automotive, medical, and clean energy. Second, in many applications, REEs are difficult to substitute with other inputs due to the high specificity of their applications (USGS, 2002), even though their input quantities are oftentimes tiny. Third, their supply over time tends to be inelastic due to their nature as byproducts as well as their toxic processing requirements (Nassar et al., 2015; EPA, 2012) and long time-to-build in mining and processing. Finally, the supply for REE mining and processing is highly concentrated. Notably, China controls approximately 60% of the mining and 90% of the post-mine processing of these elements (BIS, 2023), distinguishing them from other critical minerals whose production is geographically dispersed, such as Lithium.

<sup>&</sup>lt;sup>1</sup>Critical uses include permanent magnets, which are present in electronic devices as well as vehicle motors and wind turbines, and various chemical catalysts, which are essential for energy efficiency, environmental protection, and renewables production, USGS (2022).

An important challenge in assessing the importance of REE inputs for downstream industries is that detailed information on REE-producing and -using industries is not generally explicitly available in standard industry data and IO tables. Therefore, the first contribution that this paper makes is the construction of a novel IO table that maps the use of REEs across industries into the U.S. Bureau of Economic Analysis supply-use table. This IO table combines quantity information on the REE content of all key applications of REE at the level of individual chemical elements that we obtain from the U.S. Geological Survey with element-level price information taken from industry sources. This augmented IO table allows us to calculate total REE requirements for each 4-digit SIC manufacturing industry.

Our second contribution is an empirical investigation of an adverse REE supply shock on REE-using industries across the world. We exploit the rare earth crisis of 2010 as a quasi-natural experiment. Following a territorial dispute with Japan in 2010, China increased restrictions on its exports of REEs to the rest of the world. With China being essentially a monopolist in the extraction and sales of most REE minerals at that time, this negative REE supply shock caused a surge in international REE prices amid sourcing uncertainties. Global REE unit prices spiked by factors of 10 to 45, remaining high for about five years, as REE mine output fell and export restrictions were in place. In March 2012, the U.S. brought a case to the World Trade Organization (WTO) Dispute Settlement Body against these export restrictions on REEs. Following a WTO ruling, China relaxed its export restrictions on REEs in 2015, restoring access to these critical materials on the global market. Thus, while the shock ultimately turned out to be temporary, the duration of Chinese supply restrictions was ex ante unclear. Moreover, China's actions revealed its willingness to weaponize export restrictions, and it became clear that it could potentially repeat similar measures.<sup>2</sup>

Using data on manufacturing industries across countries, we provide evidence for the impact of Chinese export restrictions on REEs on downstream industries. Our analysis emphasizes directed technological change by investigating the impact of an REE supply shock on innovation in downstream industries, in contrast to existing literature, which typically examines the restricted inputs directly. We employ difference-in-difference estimates with continuous treatment intensity, where the exposure of each downstream industry to the contractionary REE supply shock depends on its total REE input requirement by chemical element and the substitutability of each element as determined by their physical-chemical properties (Graedel et al. 2015). To measure innovation activities downstream, we obtain the universe of granted patents across countries that mention REEs, chemical compounds of REEs, or REE-specific technologies from the Google Patent Re-

<sup>&</sup>lt;sup>2</sup>China eventually did reimpose restrictions on six REEs and rare earth permanent magnets in April 2025 as a response to the second Trump administration's new tariffs (Bradsher, 2025). This was an escalation from an earlier ban on the exports of Gallium, Germanium, Antimony and superhard materials to the U.S. in late 2024 (Pierson et al., 2024).

search Database. We use a large language model (LLM) to identify those patents within the corpus that describe key improvements of REE-using technologies and to link each individual patent to industries cross-checked with alternative methods.

Our evidence supports the hypothesis that the supply shock triggered directed technological change in REE-using industries: we find a surge of patenting activity outside of China for technologies that are using REEs and improving the efficiency of those technologies, including by substituting REEs as inputs. This surge in REE-related patents exceeds the overall growth in industry patenting. We also complement our patent analysis with panel evidence on total factor and labor productivity. While relative productivity increased in REE-intensive downstream industries outside of China in response to the adverse REE supply shock, we see relative productivity declines in China's REE-intensive industries following the REE supply shock.

Having shown that the REE supply shock triggered an increase in innovation and productivity in downstream industries, we assess the impact on the competitiveness of downstream industries by studying the evolution of exports. We find that manufacturing industries outside of China that are relatively REE-intensive expanded their exports relatively more compared to less REE-intensive industries in the same country. On average, manufacturing industries that are one standard deviation more REE-sensitive than the cross-sectoral average experienced a 0.35 percentage point (p.p.) higher annual growth rate in exports between 2010 and 2018 compared to the period 2002-2009. The effect on export growth was particularly large for manufacturing industries located in Europe or Japan. By contrast, similar export growth of REE-intensive sectors did not occur within China, where access to REEs was abundant. Aside from element-level variation, we also consider country-level variation in the exposure to the REE supply shock. Here, we consider either historical REE sourcing shares from China or country-specific spikes of REE import unit values. In line with our previous results, we find that REE-using industries outside of China expanded their REE-related patenting, productivity, and export growth relative to other industries.

Our third contribution is to quantify the impact of China's REE export restrictions on downstream industries using a novel quantitative general-equilibrium (GE) trade model featuring comparative advantage based on factor abundance and directed technological change. The model integrates a structural multi-sector gravity model of trade with a detailed IO structure (see Caliendo and Parro, 2015; Fadinger et al., 2024) with a two-factor Heckscher-Ohlin model (Chor, 2010; Burstein and Vogel, 2017; Morrow and Trefler, 2022) and a static variant of Acemoglu's (2002) model of directed technological change. In this framework, industry-level goods are differentiated by origin, and perfectly competitive firms produce output using a combination of intermediate inputs and a value-added bundle. The value-added bundle aggregates REE inputs and labor using a CES function with industry-specific substitution parameters. REE inputs are traded internationally,

with China acting as the sole supplier, thus allowing it to impose export taxes unilaterally.

We examine two model variants. In the short-run version, technology is held fixed and comparative advantage is determined by factor abundance. Under this scenario, a Chinese export tax on REEs raises input costs for downstream industries outside of China — especially when the elasticity of substitution between REE inputs and labor is low and for industries that are more dependent on REEs. This leads to a contraction of REE-intensive industries relative to labor-intensive ones in the rest of the world, while China's relative production of REE-intensive goods expands.

By contrast, in the long-run version of the model, technology is endogenous, and its factor bias is shaped by firms' targeted innovation efforts in the REE or labor input layers. A key determinant of the direction of technological change in response to a Chinese REE export tax is the elasticity of substitution between REEs and labor. If REEs and labor are gross complements, a negative supply shock to REEs shifts the direction of innovation toward REE-saving technologies. In this case, rising global REE prices increase the profitability of innovations that improve REE efficiency. When innovation externalities are sufficiently strong, this technological response can more than offset the direct cost increase resulting from higher REE prices, leading to a decline in downstream production costs of REE-intensive industries relative to labor-intensive ones. As a result, REE-intensive industries may expand relative to labor-intensive industries outside China.

We calibrate our model using detailed trade and production data from the World Input-Output Database (WIOD) for the pre-shock year 2009. Central to the calibration are the industry-level elasticities of substitution between REEs and labor, as well as each industry's reliance on REE inputs. Due to the absence of comprehensive data on REE expenditure shares, standard estimation techniques for substitution elasticities are not applicable. As a fourth contribution, we thus develop a novel structural estimation method for the elasticity of substitution between REEs and labor. This approach exploits observable differences in innovation activity — measured by the relative number of patents aimed at improving REE input efficiency — and relates them to variation in relative REE prices. Our estimates reveal that the elasticity of substitution between REE and labor is well below unity in most industries. Given these estimates, we calibrate industry-specific REE input intensities using U.S. data. We find that industries with higher REE intensity exhibit lower substitution elasticities, providing empirical support for our proposed theoretical mechanism.

In our analysis of the general-equilibrium effects of the Chinese policies on REEs, we show that the model qualitatively and quantitatively replicates the patterns observed in our reduced-form evidence. When technology adjusts endogenously, an increase in global REE prices — triggered by the export restriction and a decline in global REE supply — induces a sufficiently strong directed technological response, causing REE-intensive industries to expand relative to other industries outside of China and to contract within China. While the policy yields a modest welfare gain for

China, the negative impact on GDP and welfare in other countries is mitigated almost entirely by the endogenous technological adjustment, which significantly reduces global demand for REE. We further demonstrate that directed technological change is crucial for these outcomes. In the version of the model with fixed technologies, global production and exports in REE-intensive industries shift markedly in favor of China. In this scenario, China experiences a substantial economic boom and large welfare gains due to positive terms-of-trade effects, whereas other countries suffer considerable declines in real GDP and welfare. Overall, our analysis links the quantitative effects of export restrictions on REE to their distinctive technological features — namely low substitutability and inelastic supply—in a highly policy-relevant context.

Related Literature: Our theoretical approach builds on the literature on modern quantitative trade models. Seminal contributions include Eaton and Kortum (2002), Costinot and Rodríguez-Clare (2014), and Caliendo and Parro (2015). Caliendo and Parro (2022) provide a comprehensive overview of trade policy in this class of models, while Ossa (2014) analyzes strategic trade policy, and Lashkaripour and Lugovskyy (2023) examine the interaction between trade and industrial policies in settings with scale economies. We contribute to this literature by incorporating directed technological change into a quantitative trade model. Although the existing literature largely examines import tariffs, our focus is on export taxes and quantity restrictions.

In the field of innovation, our work relates to studies such as Acemoglu (1998, 2002), Acemoglu et al. (2012), and Aghion et al. (2016) on directed technological change. The latter authors find that firms in the auto industry innovate relatively more in clean technologies when they face higher tax-inclusive fuel prices but document path dependence in the type of innovation (firms' history and aggregate spillovers). In terms of a global economy, Acemoglu et al. (2015) highlight the complex relationship between direction of technological progress and offshoring, which can both drive skill-biased technological change and spur innovations favoring unskilled labor. More generally, a series of papers on biased technological change, document how innovation favors a particular input due to its relative supply or price, with Kennedy (1964) being one seminal work. In a historical context, Hanlon (2015) finds that the blockade on cotton from the Southern U.S. during the U.S. Civil War spurred technological progress in the use of cotton inputs from India in a context where inputs are gross substitutes. Popp (2002) finds that higher energy prices induced more innovation in energy-saving technology, proxied by patents. Hassler et al. (2021) model energy-saving technological change to estimate the elasticity of substitution between energy and labor or capital inputs. Relatedly, Blum (2010) documents empirically for a large sample of countries that changes in countries' relative factor endowments do not only lead to Heckscher-Ohlin forces by shifting the output mix of economies' but also change factor returns in the long run, which leads to directed technological change, shifting producers' isoquants in the long run.

In contrast to existing work, we emphasize the propagation effects of a deliberate industrial policy intervention targeting an essential and concentrated input—characterized by low substitutability and inelastic supply on downstream innovation and productivity in third countries in a modern setting. Furthermore, our quantitative GE analysis of REE tax export restrictions provides a detailed welfare assessment of their impact on downstream innovation.

The academic literature on industrial policy is vast, as summarized in Harrison and Rodríguez-Clare (2010) and Juhász et al. (2024). Empirical evidence on the effect of industrial policy is mixed, fueling the debate. Criscuolo et al. (2019), for example, find positive effects on incumbent firms' investment and employment but not TFP for regional policies. Bartelme et al. (2025) explore the theoretical and empirical evidence of industrial policy subsidies based on external economies of scale, finding limited empirical support. The analysis reveals significant variation in economies of scale across manufacturing sectors. However, in highly open economies, the impact of such policies appears to be minimal and not transformative. Studying industrial policy in high-tech sectors, Barwick et al. (2024) find that targeted subsidies increase firm-level innovation but have modest spillover effects on industry productivity. The authors highlight risks of resource misallocation, questioning the efficiency of such interventions. Kee and Xie (2025) examine Indonesia's export restrictions on nickel and highlight the unintended negative impact on Indonesia's own downstream industries as lower domestic nickel prices enable the entry of smaller, less efficient steel-using firms. In contrast, Juhász et al. (2024) offer a more nuanced and generally positive perspective on industrial policy, highlighting its potential to drive structural economic transformation.

Equally rich is the literature on supply chains and trade restrictions, with a recent overview by Fajgelbaum and Khandelwal (2022) and significant works by Grossman et al. (2024), and Bown et al. (2023). A vast number of papers document the effects of trade restrictions on global supply chains. Liu (2019) finds that targeted industrial policies can enhance efficiency by correcting distortions in upstream sectors. Barattieri and Cacciatore (2023) show that protectionist trade barriers disrupt production networks, harming downstream firms through higher input costs and job losses. Recent studies focusing on U.S. tariffs document a "great reallocation" of import sourcing away from China, documenting widespread reallocation and negative effects on prices and welfare (Amiti et al. 2019, Fajgelbaum et al. 2020, Flaaen et al. 2020, Grossman et al. 2024, Alfaro and Chor 2023, Alfaro et al. 2025).

The paper also relates to the growing literature on the role of sanctions and geoeconomics more broadly (Hirschman, 1945; Felbermayr et al., 2020). A large portion of this literature has analyzed financial effects (e.g., Cipriani et al., 2023; Eichengreen et al., 2023; Itskhoki and Mukhin, 2022), while our paper focuses on downstream production and innovation implications. Juhász (2018), for example, finds increased mechanized cotton spinning in French regions more exposed to British

Blockage. Moll et al. (2023) highlight the different ways in which German firms and households adopted to the restricted imports of energy through demand reduction, increase efficiency, and diversification to alternative sources. Our analysis stresses and quantifies the role of directed technological change in adjusting to adverse shocks in a modern context and highlights that in an environment of strongly complementary inputs sanctions may backfire by triggering foreign innovation.

The remainder of the paper is organized as follows. Section 2 describes the context and policies around REEs. Section 3 presents the data and describes the construction of the IO table and the REE-related patent data. Section 4 presents the empirical analysis of innovation and trade effects. We then present the model and its quantification in Section 5 and the last section concludes.

# 2 Background: Rare Earth Elements

## 2.1 Differentiating Characteristics

The U.S. Geological Survey (USGS) defines the REEs as a group of 17 elements composed of Scandium, Yttrium, and the lanthanides.<sup>3</sup> We note four characteristics of REEs that distinguish them from other minerals.

First, these elements are collectively known for their unique magnetic, catalytic, and luminescent properties, making them *essential* in a broad variety of high-tech and strategic applications (USGS, 2014). These include electronics, lighting, aerospace, defense, medical and green technologies. These diverse applications arise from chemical properties that are similar across all REEs, hence their classification as a group, although slight differences in electronic configurations give individual elements specific specializations.<sup>4</sup>

Second, their applications usually involve small quantities but they are challenging to substitute due to their high specificity (USGS, 2002; Graedel et al., 2015). For example, no other known elements could replace Europium as a red phosphor for monitors or Erbium in laser repeaters for fiber optics. Meanwhile, in principle, REE permanent magnets in electric vehicle motors can be substituted with ferrite magnets, but this would make them about 30% heavier (Adamas

<sup>&</sup>lt;sup>3</sup>The lanthanides include the following elements: Lanthanum (La), Cerium (Ce), Praseodymium (Pr), Neodymium (Nd), Promethium (Pm), Samarium (Sm), Europium (Eu), Gadolinium (Gd), Terbium (Tb), Dysprosium (Dy), Holmium (Ho), Erbium (Er), Thulium (Tm), Ytterbium (Yb), and Lutetium (Lu).

<sup>&</sup>lt;sup>4</sup>Salient examples of REE use are for permanent magnets (e.g., Neodymium and Samarium), as found in products ranging from electronic speakers and medical equipment to wind turbines and catalysts (e.g., Cerium, Lanthanum) used in petroleum refining and automobile exhausts and medical treatments (e.g. Gadolinium for magnetic resonance imaging, Yttrium for radiation therapy). See Voncken (2016) for a list of applications by element as related to their properties.

Intelligence, 2023). This characteristic contributes to their *low elasticity of substitution* as an input, which informs economic responses to shocks on their supply.

Third, REEs have a relatively *inelastic supply* due to their nature as byproducts (Nassar et al., 2015) along with their toxic processing and handling (EPA, 2012). Despite their name, most REEs are geologically relatively abundant. However, due to their high reactivity, REEs do not occur as individual metals in nature the way copper and silver do, but instead as constituents of ores and minerals (Balaram, 2019). Lanthanum, La ('lanthanein,' lying hidden), for example, occurs with Cerium, Ce. The deposits were found later to contain also Praseodymium ('prasinos,' leek-green, and 'didymos,' twin) and Neodymium ('neos,' new twin). High similarity in their chemical configuration, or their "chemical coherence", makes them particularly difficult to separate from each other (The Geological Society of London, 2011). Consequently, mining and processing REEs involve steps that are relatively lengthy, complex, and costly (Hurst 2010). Other chemical properties, such as radioactivity, complicate the process and add regulatory costs.<sup>5</sup>

Finally, REE production and processing are one of the most concentrated across mineral resources, with China controlling more than 90% of post-mining processing (BIS, 2023). China possesses abundant REE resources, including the only developed ion-adsorption clay deposits, which are the most low-cost source for heavy REEs (Packey and Kingsnorth, 2016; USGS, 2002). This natural endowment was compounded by China's long history of REE promotion policies.

### 2.2 China's REE Policies and Production

In 1975, China sought to promote the REE mining industry (Shen et al., 2020). Policy tools included export-tax reimbursements starting in 1985, which coincided with China's REE market share growth. In 1990, the Chinese government designated REEs as a strategic resource, barring foreign investors from ownership of mines and limiting their involvement to REE smelting and separation projects unless through joint ventures with Chinese companies. Every REE mining or smelting project and joint venture required approval from the state, and additional export quotas were introduced in 1999 to address illegal production (Tse, 2011).

In the early 2000s, China emerged as the dominant player in the global production and processing of REEs, driven by several factors including abundant reserves, low labor costs, lax environmental

<sup>&</sup>lt;sup>5</sup>Most REE mines have Thorium, for example, which is radioactive. Cerium, Lanthanum, and Dysprosium, like the other lanthanides, react easily with Oxygen (corroding quickly), are highly reactive to water, highly pyrophoric (igniting spontaneously in the air), and are powerful irritants requiring particular handling protocols. REEs have substantial metabolic effects: skin exposure can result in irritation, ulceration, delayed healing, and granuloma formation; ocular contact can lead to conjunctivitis and corneal damage, scarring, and opacity. Inhalation of REE dust can induce acute irritative bronchitis and pneumonitis. Most REEs are classified as mild to moderate toxic (Harbison and Johnson 2015).

regulations, and economies of scale. Not long before, REE processing plants in advanced economies such as France and Japan began shutting down, culminating in the closure of California's Mountain Pass, the primary source of REEs before the mid-1980s (USGS, 2002). Between 1990 and 2000, China's production grew by more than 450%, from 16,000 to 73,000 metric tons (Tse 2011). During this same period, production from other countries dropped by nearly 60%, from 44,000 to 16,000 metric tons. As a result, global production grew by just over 150% (from 60,000t to 91,000t). China's share of global REE output surged from 27% in 1990 to over 90% by 2008, supplying the vast majority of the world's REEs in the form of concentrates, intermediate products, and chemicals during this period. In 2009, China's market share in global REE mine production reached 98% (see Figure A.1 in the Data Appendix).

# 2.3 The 2010 REE Supply Shock

We exploit the extreme and unprecedented 2010-2011 REE price surge by a factor of 10 (Terbium and Europium) to 45 (Cerium) as a negative supply shock for our empirical analysis (see Figure 1).<sup>6</sup> The price jump was triggered by new Chinese export restrictions that tightened supply and heightened uncertainty in the context of geopolitical tensions. Beyond the extreme price spike, the main impact of China's export restrictions was supply uncertainty. The episode demonstrated China's ability—and willingness—to weaponize REEs. Concerns over China's near-monopoly on REEs had been minimal before 2010, as reflected in the scarcity of official reports and business discussions.

The 2010 crisis made the risk of future disruptions evident, creating a lasting shock to REE supply uncertainty, which we focus on in our analysis. In July 2010, China drastically reduced the REE export quota for the second half of the year by 72%, with a stated motive of combating illegal mining and environmental pollution (Müller et al., 2016).<sup>7</sup> Then on September 7, 2010, a Chinese trawler collided with Japanese coast guard boats in the Senkaku-Diaoyu waters, triggering a diplomatic dispute. Immediate reports emerged of a ban as China halted all shipments of REE in retaliation for Japan's detention of a Chinese fishing boat captain. The restrictions were soon followed by increased export tariffs on certain REE products in January 2011, and higher taxes on REE mining and new export quotas on REE ferroalloys by May 2011 (WTO 2012, OECD 2024). REE market prices jumped at an unprecedented rate, peaking in mid to late 2011 before gradually

 $<sup>^6</sup>$ REE markets are highly illiquid, making yearly price changes more discernible than higher-frequency movements, which tend to be extremely noisy. For comparison, the factor price increase from the 1974 OPEC shock was around 300%, while the REE shock was between 1000% and 4500%.

<sup>&</sup>lt;sup>7</sup>In August 2009, China had issued a draft policy to cut the annual export quota with potential export bans on certain heavy REEs and reduced annual REE export quotas by around 12% each year from 2005 to 2009 (see Shen et al. 2020; Packey and Kingsnorth 2016; Pritchard 2009; Bradsher 2009, 2010). The domestic REE production quota was also cut by 20% in 2010 (Tse, 2011).

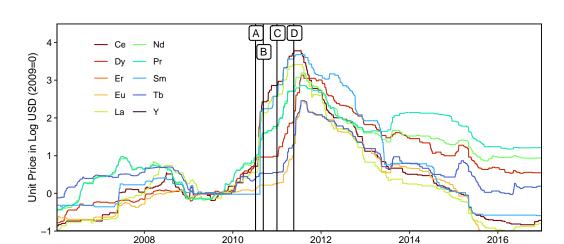


Figure 1: Unit Prices of Selected Rare Earths

Notes: The figure plots indices of REE log unit prices using data from Asian Metal. The select elements are those with complete requirement data in the USGS report (Bleiwas and Gambogi, 2013). The vertical lines denote relevant episodes: A=Export quota cut by China; B=Senkaku-Diayou boat collision; C=Select export tariffs hike by China; D=New export quotas on ferro-alloys by China.

2014

2016

declining. Global REE supply dropped by about 30,000 mt as China tightened production (Figure A.1).

On March 13, 2012, the U.S., joined by Japan, the E.U. and Canada, requested WTO dispute settlement consultations. A WTO panel dismissed China's exemption argument based on conservation of an exhaustible natural resource and determined that the quotas were intended for industrial policy objectives (WTO 2015) and in 2014 ruled in the plaintiff's favor. China removed its export quotas on REEs in 2015, replacing the quota with an export license and the export tax with a resource tax based on value (Mancheri et al., 2019).

**Supply Chain and Innovation Responses:** The price spike spurred efforts to explore alternative sources of REEs, including new mines and processing facilities. Progress was slow and limited, hinting at the low supply inelasticity of REEs. Mining expanded in Australia, while firms in the U.S. continued previously ongoing efforts to reopen Mountain Pass, though it took at least two years for new supply to reach markets, and both mines operated below planned capacity.<sup>8</sup> Production and processing remain reliant on China. Uncertainty constrained inventory

<sup>&</sup>lt;sup>8</sup>Mountain Pass, sold to Molycorp in 2009, reopened in 2015 but soon shut down due to bankruptcy. It resumed operations in 2018 under MP Materials, with Chinese REE miner Leshan Shenghe holding a non-voting minority stake (Topf, 2017; Brickley, 2017). Lynas's plant in Malaysia commenced light REE oxides production in 2013 after the Malaysian government approved its development in 2007 (Gholz, 2014; Lynas, 2007).

use. Many Japanese firms, for example, stockpiled rather than released REEs, fearing depletion without fresh imports from China (Gholz and Hughes, 2021). Firms sought to find substitutes for REEs, alternative manufacturing processes, improve efficiency and recycling efforts. Given the small quantities of REEs used in many applications, large-scale recycling was economically nonviable (Hurst, 2011). In addition, firms, from auto companies in Japan, such as Toyota and Honda, and abroad, such as Tesla, Renault, GM, and electronic companies, such as Hitachi and Phillips, announced efforts as soon as 2010 to create new products using less or no REEs (Banner, 2022; Reuters, 2021; Bomgardner, 2018; Owano, 2018; Halvorson, 2022; Houser, 2023). 10

# 3 Data Sources and Descriptive Statistics

## 3.1 Input-Output Table with Rare Earths

To construct an IO table with REEs, we combine data from different sources. In particular, we capture the use of REEs from USGS and prices from BCC and Asian Metal. In Appendix A.1, we provide details on the imputation process of REE use into the standard supply-use table.

Use of REEs: We start with the 2012 supply-use table from the Bureau of Economic Analysis, the closest available to the time of the supply shock. The supply-use table comprises 405 BEA industry groups, which are a slightly more aggregated version of the 6-digit 2012 NAICS codes. This table reports the value of inorganic chemicals used by each industry—a broad category that includes REEs. To zoom in on REEs, we augment this table with data from a USGS report on REE inputs consumed by the U.S. in 2010 (Bleiwas and Gambogi, 2013). This report combines secondary data on REE inputs from various sources as well as primary data from disassembling and analyzing the manufacturer labels of numerous products. Total use of each element is computed by multiplying the amount of each REE contained in a product by the total quantity of the product consumed in the U.S. as obtained from consumption and trade statistics. It is the most comprehensive REE content data available for the time around the supply shock episode. REE use is available for seven "General Application" categories, namely alloys, batteries, automobile

<sup>&</sup>lt;sup>9</sup>As noted by Bachmann et al. (2024), the actual response by firms and governments is consistent with the embargo's effectiveness in creating uncertainty on the reliability of the supply chain.

<sup>&</sup>lt;sup>10</sup>In January 2011, GM filed a patent for a powder coating process that reduced Dy and Tb usage in REE magnets by at least 20% preserving their magnetic properties, noting supply constraints in the filing. In July 2011, Skyworks Solutions patented a yttrium substitute in synthetic garnets for electronic microwave devices, citing rising costs due to restricted Yttrium supply. By 2016, Toyota cut Dy in the Prius and, in 2018, reduced Nd in electric motor magnets by 20%. Volkswagen followed, while Nissan and BMW unveiled magnet-free motor prototypes in 2022-2023. Other industries also sought REE reductions, such as catalysts using less Cerium (Machida et al., 2017). See Appendix B.3 for other examples.

catalysts, fluid catalyst cracking, phosphors and diodes, and solutions.<sup>11</sup> These categories are the most upstream available for our purpose. We focus on estimates of the use of five broadly used REEs: Lanthanum, Cerium, Praseodymium, Neodymium, and Dysprosium.

REE Price Data and Total Requirements: As the estimates are in metric tons, we convert them into USD million using a combination of prices from BCC (2015) and Asian Metal, both at the element level. Finally, we use the supply-use table, augmented with imputed REE data, to compute the total requirements matrix (Leontief inverse). The entries of this matrix indicate the amount of input each industry receives from every other industry, accounting for both direct and indirect linkages. Table A.4 in the Data Appendix lists the most REE-intensive industries. The top using industry by total requirement is SIC 3691 (Storage batteries), which uses Lanthanum. Another REE-intensive industry is SIC 3625 (Relays and Industrial Controls), which employs Neodymium, Praseodymium, and Dysprosium.

### 3.2 Patent Data

Our primary measure for shifts in the direction of technological change is based on patent data. We obtain the universe of granted patents related to REEs from the Google Patent Research database.

REE Patents: As a first step, we identify patents as broadly related to REEs and link them to individual elements when their title or abstract contains specific keywords that include either the name of the elements themselves, their chemical compounds, or some key related technologies, such as technologies related to permanent magnets (Table A.5 in the Data Appendix lists the keywords). By the end of the sample period, around 30,000 distinct REE patents had been granted worldwide. We assign each patent to a country based on the location of the patent assignee, focusing on the top 50 countries by GDP. These countries are grouped into the following regions: Europe, the United States and Canada, China, Russia, South Korea, Japan, Australia, and the Rest of the World. Table A.1 in the Data Appendix lists the countries in the sample. In the next step, we link each patent in our sample to a corresponding industry. This task is non-trivial, as patents are classified by technological fields rather than by industries. Given the substantial share of patents filed by non-corporate entities and the absence of a comprehensive global patent—firm matching database, we employ a large language model (LLM) to assign patents in our sample to industries.

<sup>&</sup>lt;sup>11</sup>We exclude phosphors and diodes due to the presence of Europium and Terbium, whose estimates in dollar value become highly uncertain when multiplied by their high per-unit prices, which are one to two orders of magnitudes larger than for the rest of the elements. We also exclude solutions because the report provides no precise point estimates for this category's input element use.

Specifically, we extract and parse the title and abstract of each patent, and prompt the LLM (GPT4 from OpenAI) to identify the SIC code that best corresponds to the described technology.

Furthermore, we refine the sample of REE-related patents by instructing the LLM to classify each patent according to whether it pertains to a technology that enhances the efficiency of using REEs or facilitates substitution away from REEs.<sup>12</sup>

As a consistency check of our patent classification, we regress the cross-section of patent stocks measured at the industry-element level on its total requirement share from our IO table, including a full set of SIC industry fixed effects. An industry with a REE requirement for a specific element that is one standard deviation above the sample mean has, on average, 8.5% more patents mentioning this REE. We present examples of such patents in Appendix A.8.

Non REE Patents: One concern with our approach to measuring directed technological change is that we might pick up overall innovation in the industry and not necessarily innovation that is directed explicitly towards increasing the efficiency of REEs. We address this concern by controlling for the overall stock of patents in a given region-industry. Since the number of patents by region-industry is not directly observable, we proceed as follows. First, we draw a large random sample of patents and let the LLM allocate them to SIC industry codes. We then scale these numbers across regions by the number of granted patents that each region has when considering the full sample of manufacturing patents for that year.

# 3.3 Trade and Industry Data

We use data on exports of 4-digit SIC level manufacturing industries for the 50 largest economies in the world from 2002 until 2018 from UN Comtrade. We build a country-industry panel of total factor productivity and labor productivity to measure the effect on productivity across countries and industries. For Japan and the U.S., we have productivity measures from the Japanese Manufacturing Census and the NBER CES manufacturing database, respectively. For the other countries, we rely on data from the UNIDO Indstat database. Appendix A.4 provides details on the productivity data. To control for industry characteristics, we obtain the average capital and labor intensity from the NBER CES manufacturing database.

We construct a country-industry proxy of how strongly an industry is targeted by industrial policy using the subsidy database from the Global Trade Alert (GTA). This proxy is the share of subsidies

<sup>&</sup>lt;sup>12</sup>As an illustrative example, consider U.S. patent US-8586678-B2 with the title "Blends of linear and branched neodymium-catalyzed rubber formulations for use in golf balls." The LLM. classifies this patented technology as improving the usage of REEs and assigns it to SIC 3949 "Sporting and Athletic Goods, n.e.c." which includes golf equipment.

- counted as the number of policy measures - going to a given country-industry relative to the total number of subsidy policies worldwide.

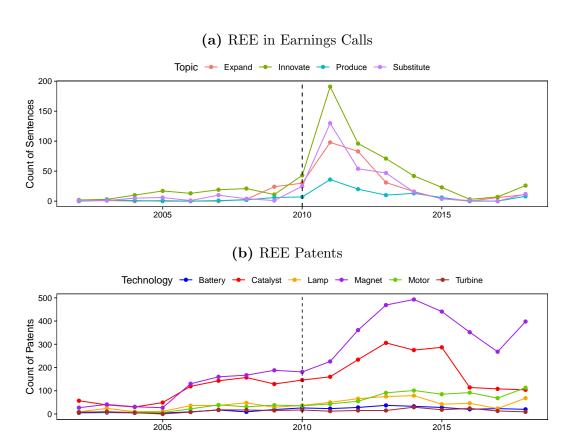
## 3.4 Descriptive Patterns

Figure 2 shows composite REE price indices (log USD) over the sample period. Prices surged around 2010, peaked shortly after, then declined sharply and stabilized at a lower level. At the individual REE level, prices rose by a factor of 10 to 45. This spike reflects the global supply shock triggered by Chinese policy interventions in the REE market. Although prices did not remain elevated in the long run, their subsequent decline was driven by a medium-term increase in supply—primarily resulting from the relaxation of Chinese export restrictions and, to a lesser extent, limited entry by foreign producers—as well as a reduction in demand attributable to innovation. As we show in the empirical analysis below, the strong innovation response suggests that downstream industries anticipated that China's export restrictions might be permanent or might recur in the future.

Figure 2 depicts additional descriptive patterns present in the data. Panel (a) highlights how firms communicate about the negative REE supply shock. We scrape earnings calls of international firms and count sentences that mention REEs or related keywords. We then use an LLM to categorize those sentences. "Expand" denotes sentences about the expansion of REE production (including mining and refining), "Innovate" denotes innovations in REEs, which are further subdivided into production of REEs ("Produce") or usage substitution ("Substitute"). The mentions rise sharply around the 2010 price spike, particularly for innovation-related topics. Within the innovation category, most mentions discuss substitution, highlighting a heightened focus on exploring alternatives in response to the price shock.

Panel (b) depicts counts of newly granted patents that relate to REE. These patents generally increase over time, particularly in magnet and catalyst technologies, indicating a growing emphasis on innovation in these areas. However, the trend strongly accelerates after 2010, in line with the idea that the price spike may have catalyzed greater investment in technological development and innovation related to REEs. Overall, Figure 2 suggests that the sharp increase in REE prices around 2010 significantly influenced producer behavior and innovation activities. Finally, the surge in patenting is hump-shaped and disappears after 2015, once Chinese export restrictions have been lifted. This indicates that it was not driven by a permanent sectoral shift towards green technologies, such as electric vehicles and wind turbines, but rather by a temporary price hike of REEs.

Figure 2: The REE Supply Shock, Earnings Calls, and Innovation



Notes: The figure presents mentions of REEs in firms' earnings calls (a), and patents related to REEs (b). Subfigure (a) depicts the count of REE mentions in global firms' earnings calls, categorized by topic using an LLM. Topics include "Expand" (expansion/diversification of REE production) and "Innovation" (innovations in REE production or usage). Innovation mentions are further split into "Produce" (innovating REE production) and "Substitute" (reducing/substituting REE use). Data is sourced from NL Analytics. Subfigure (b) shows the yearly count of new patents granted related to REE technologies, identified through keyword analysis in patent abstracts. Data is sourced from Google Patents.

# 4 Empirical Analysis

## 4.1 Directed Technological Change

Innovation in Downstream REE-Using Industries: As a starting point, we explore whether the negative REE supply shock had an impact on the direction of technological change in downstream manufacturing industries across the world and in China, using patent data. We consider a differences-in differences specification with variable treatment intensity. We estimate the following count model for patents:

$$y_{rst} = \beta REE \ Sensitivity_s \times post_t + \gamma \Delta_{rst} + \eta_{rs} + \eta_{rt} + \epsilon_{rst}. \tag{1}$$

The outcome variable  $y_{rst}$  is the stock of granted REE patents for the 4-digit SIC manufacturing industry s from region r in year t, considering the sample window from 2002 until 2018. Our coefficient of interest  $\beta$  is the coefficient on an interaction term of the industry's REE sensitivity measure REE Sensitivity, described below with a step dummy  $post_t$  which is one from the year 2010 onward. This corresponds to the period after China first implemented export restrictions and proxies for regime change in the supply-chain environment related to REEs.

We include a full set of region-by-industry and region-by-year fixed effects that control for the average level of patenting at the region-industry level and region-specific shocks, respectively, as well as a set of industry-region-period-specific control variables that we describe below. Note that the inclusion of region by industry fixed effects implies that the identification is based on time variation in patent stocks within a given industry-region pair, so that we effectively investigate the impact of the treatment on patent flows.

As a measure for the exposure of an industry s to the REE supply shock, we make use of the element-level total requirements from our constructed IO table and an index of element-level REE complementarity from Graedel et al. (2015). This complementarity stems from the physical and chemical properties of elements, not economic factors, and is based on expert assessment. Dysprosium, for example, ranks as the least substitutable, while Samarium is the most. Between two industries using equal REE amounts, the one relying on harder-to-substitute elements faces greater exposure to REE supply shocks. For instance, industries needing Dysprosium for permanent magnet motors—with inadequate substitutes— are more exposed than those using REEs in nickel-metal hydride batteries, which can be replaced by lithium-ion alternatives.

To construct the exposure variable  $REE\ Sensitivity_s$ , we multiply total requirements  $tr_{es}$  of each REE e for industry s with the index of complementarity  $compl_e$  (ranging between 0 if perfect

substitutes are available to 100 if an element cannot be substituted at all) and aggregate over the REE inputs:

$$REE \ Sensitivity_s = \sum_e tr_{es} \times compl_e. \tag{2}$$

This variable measures the exposure of industry s to the REE supply shock. REE Sensitivity<sub>s</sub> is high if the total requirements of REE inputs are large and these REE inputs are very complementary.

We estimate (1) using Poisson pseudo-maximum likelihood estimation. An important control variable that we include throughout all specifications, is the total patent stock in industry s from region r during year t (in logs) to proxy for overall movements in the industry's innovation activity that might be unrelated to innovation in REE-related activities. This alleviates concerns that our estimates may pick up region-industry-specific trends, such as a general surge in patenting in industries that are affected by the green transformation. A further related concern is that the surge in economic activity of REE-intensive industries may be driven by local subsidies or an increase in demand for their output (e.g., because some of REE-intensive industries are important for the green transformation). We thus additionally control for the capital and labor intensity of each industry (time-invariant) as well as for the (time-invariant) industry-region-specific subsidy intensity (from the Global Trade Alert database), all interacted with the treatment-period dummy  $post_t$ . Furthermore, we control for a time-varying country-industry-specific demand proxy. This demand proxy is the lagged weighted GDP growth of the top 10 importing countries in year t (measured in midpoint growth rates) by industry-region. We cluster standard errors at the country-industry level, following Abadie et al. (2023).

Table 1 reports our estimates of (1). The sample only counts those patents that the LLM identifies as patents that either improve the efficiency of REEs or help to substitute the use of REEs. We find that industries that are more sensitive to the REE supply shock innovated more in the treatment period than before compared to less affected industries. For the non-China subsample, a one standard-deviation higher value of REE sensitivity is associated with 7.4% larger stock of REE-enhancing patents after the 2010 supply shock. This effect was particularly strong in European countries, the U.S. and Japan. Perhaps not surprisingly, industries more sensitive to REE inputs also patented slightly more REE technologies in China during the treatment period compared to industries. However, this difference in patenting behavior was much less pronounced than in other economies and is statistically indistinguishable from zero.

**Productivity Effects:** Next, we study the impact of the policy change on measures of productivity of REE-using industries. We estimate the following difference-in-differences specification,

Table 1: Patents in Rare-Earth Intense Manufacturing Industries

	REE-Enhancing Patents					
	NONCHN (1)	ALL (2)	USA (3)	EUR (4)	JPN (5)	CHN (6)
REE Sens. $\times$ Post	15.46***	14.36***	17.91**	18.43***	25.65**	2.886
	(4.943)	(5.155)	(7.227)	(6.026)	(11.05)	(19.35)
Observations	5,561	7,606	1,200	1,140	972	2,045
Clusters	387	531	81	74	66	144
	Yes	Yes	Yes	Yes	Yes	Yes
	Yes	Yes	Yes	Yes	Yes	Yes
	Yes	Yes	Yes	Yes	Yes	Yes

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{rst} = \beta$  REE Sens.  $\times$  post,  $+ \gamma \Delta_{rst} + \eta_{rs} + \eta_{rt} + \epsilon_{rst}$  with Poisson pseudo-maximum likelihood estimation. The outcome  $y_{rst}$  represents granted REE-related patents that improve the efficiency of REE or help find ways to substitute REE usage. The sample includes 4-digit SIC manufacturing industries (with at least one REE-related patent) from 2002-2018 across 8 regions. Regions capture the location of the patent assignee and include Australia, China, European Union, Korea, Russia, Japan, U.S. and the Rest of the World. The treatment intensity REE Sens. is a weighted sum of an REE element-specific complementarity index (ranging from 0 to 100), with weights based on total requirement shares for industry s: REE Sens. =  $\sum_e tr_{es} \times compl_e$ . post, is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Region subsamples include non-China, all regions, the U.S., European economies, Japan, and China. All regressions include region-industry and region-year fixed effects. The control vector  $\Delta_{rst}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and subsidy fractions from the Global Trade Alert database, all interacted with post, as well as the log of the total stock of granted patents in region-industry rs in year t. It also includes a lagged demand control at the region-industry-year level, constructed by taking the log of the weighted yearly real GDP of the top 10 importer countries for that region-industry. The top 10 importer countries are identified by ranking importer countries by trade value in the period 1996-2009. In the cases of USA and CHN, this demand control is excluded due to high collinearity with the fixed effects. Standard errors (in parentheses) are clustered at the region-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\*\* p < 0.01.

using productivity growth as the outcome variable:

$$y_{ist} = \beta REE \ Sensitivity_s \times post_t + \gamma \Delta_{ist} + \eta_{is} + \eta_{it} + \epsilon_{ist}. \tag{3}$$

The outcome variable  $y_{ist}$  is the annualized growth rate of productivity for the 4-digit SIC manufacturing industry s by country i during year t, again considering the sample window from 2002 until 2018. The annualized productivity growth rate for the period t is computed using the midpoint of t and t-1 as the denominator. The coefficient of interest  $\beta$  is again the coefficient on an interaction term of our REE Sensitivity, measure with the treatment dummy  $post_t$ . All estimations include a full set of country-industry and country-year fixed effects  $\eta_{is}$  and  $\eta_{it}$ . We also include the time-varying industry controls described above (demand, subsidies, industry characteristics), weight regressions using export weights<sup>13</sup> and cluster standard errors at the country-industry level.

The regression results presented in Table 2 highlight the impact of the REE supply shock on productivity growth of REE-using manufacturing industries across different country groups. The top panel focuses on total factor productivity (TFP) growth, while the bottom panel examines labor productivity growth. In the full sample (column 2), there is a positive and significant impact of REE shock exposure on both TFP and labor productivity growth. A one standard-deviation higher value in REE sensitivity is associated with 0.19 and 0.16 percentage-point higher growth rate of TFP and labor productivity, respectively, after the REE supply shock. More exposed European and Japanese industries in particular experience a significant increase in TFP (columns 4 and 5). Japanese industries with a one standard-deviation higher REE sensitivity boast around 0.5 percentage-point higher growth rate in productivity, both using the TFP and labor productivity measures. In contrast, Chinese industries (columns 6 and 12) show a negative response, with a significant decline in TFP growth and an insignificant but negative coefficient for labor productivity growth. These results suggest that while REE-using industries in some economies, particularly in Europe and Japan, adapted to the Chinese REE policy with productivity gains, Chinese industries themselves faced relative productivity losses.

#### 4.2 Trade Effects

Effects on Downstream Manufacturing Exports: Next, we study how manufacturing exports from third countries and from China itself respond to the Chinese REE export restrictions, using 4-digit SIC trade data from UN Comtrade for our set of sample countries over the period

<sup>&</sup>lt;sup>13</sup>We use the share of export values  $w_{is} = exports_{is}/\sum_{i,s} exports_{is}$  in the period 2002-2004 as weights.

<sup>&</sup>lt;sup>14</sup>European economies in our sample of 50 countries include Austria, Belgium, Switzerland, Czech Republic, Germany, Denmark, Spain, Finland, France, United Kingdom, Ireland, Italy, Netherlands, Norway, Poland, Portugal, Romania, and Sweden.

Table 2: Productivity Growth of Rare-Earth Intense Manufacturing Industries

	An	nualized G	rowth: To	otal Factor	r Productiv	rity	
	NONCHN	ALL	USA	EUR	JPN	CHN	
	(1)	(2)	(3)	(4)	(5)	(6)	
REE Sens. $\times$ Post	0.498* (0.260)	0.466* (0.259)	0.479* (0.281)	0.555** (0.242)	1.029*** (0.370)	-2.142*** (0.742)	
Observations Clusters	183,818 14,981	186,770 $15,350$	6,323 $452$	88,055 $6,306$	5,679 414	$2,952 \\ 369$	
	Annualized Growth: Labor Productivity						
	NONCHN	ALL	USA	EUR	JPN	CHN	
	(7)	(8)	(9)	(10)	(11)	(12)	
REE Sens. $\times$ Post	0.391** (0.176)	0.377** (0.175)	0.0825 $(0.565)$	0.202 (0.159)	1.039** (0.417)	-0.633 (0.764)	
Observations Clusters	183,818 14,981	186,770 15,350	6,323 452	88,055 6,306	6,482 436	2,952 369	
	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{ist} = \beta$  REE Sens.  $_s \times post_t + \gamma \Delta_{ist} + \eta_{is} + \epsilon_{ist}$ , where  $y_{ist}$  represents the annualized growth rate of productivity for country-industry is in year t, TFP (upper panel) or labor productivity measured as value added per worker (lower panel). The annualized growth is calculated using the midpoint between t and t-1 as the denominator. The sample includes 4-digit SIC manufacturing industries from 2002-2018 across the 50 largest economies. The treatment intensity REE Sens.  $_s$  is a weighted sum of an REE element-specific complementarity index (ranging from 0 to 100), with weights based on total requirement shares for industry s: REE Sens.  $_s = \sum_e tr_{es} \times compl_e$ .  $post_t$  is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, Japan, and China. For the country subsamples of the U.S. and Japan, we use the data from NBER-CES Manufacturing Database and Japan's Annual Manufacturing Census, respectively (see Appendix A.4 for notes). All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and subsidy fractions from the Global Trade Alert database, all interacted with  $post_t$ , and the lagged weighted average growth rate of GDP of the ten largest importers from is. Regressions are weighted by the share of export values  $w_{is} = exports_{is}/\sum_{i,s} exports_{is}$  from 2002-2004. Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

from 2002 until 2018.

We again estimate (3) on various sub-samples and alternatively include all countries except China, all countries, the U.S., European economies, Japan and China, using export growth as outcome variable. All estimations include a full set of country-industry and country-year fixed effects  $\eta_{is}$  and  $\eta_{it}$  and the set of time-varying industry controls (demand, subsidies, industry characteristics). Estimated standard errors are again corrected for clustering at the country-industry level and we weight by the start-of-sample export shares.

The upper panel of Table 3 presents the results from specification (3) using export growth as the outcome variable. The coefficient  $\beta$  on the interaction term  $REE\ Sensitivity_s \times post_t$  is positive and statistically significant at the 1% level in column (1) for the sample of countries excluding China. This indicates that, outside China, exports of industries more reliant on REE inputs grew significantly faster relative to less sensitive sectors within the same country during the treatment period compared to before. Quantitatively, a one-standard-deviation higher value in REE sensitivity is associated with a 0.35 percentage point larger midpoint growth rate of exports. The coefficient remains positive across all samples except for the case of China, where it is negative but not statistically significant. The effects are particularly pronounced for European countries and for Japan, while they are more moderate in the U.S.<sup>15</sup>

The bottom panel of Table 3 presents results from specifications where the outcome variable is defined as the difference between the midpoint export growth rate of a given country-industry outside China and the one of the corresponding industry in China, thereby differencing out any industry-specific shocks. We find that, relative to the same Chinese manufacturing industries, exports of more exposed industries outside of China grew significantly more in response to the REE supply shock.

To make sure that our results are not driven by pre-trends in REE-intensive industries compared to other industries, Figure 3 visualizes the total exports of high REE-sensitivity industries versus all other industries. The two trends were parallel before the event in 2010 aside from the slight shift around the trade collapse of the Global Financial Crisis. Moreover, in line with the results in Table 3, we do observe an increase in REE-intensive manufacturing exports relative to other exports starting after the Chinese REE exports restrictions. Note that, consistent with the innovation channel, the relative increase in exports takes some time build up and starts to decline again after sanctions are lifted in 2015.

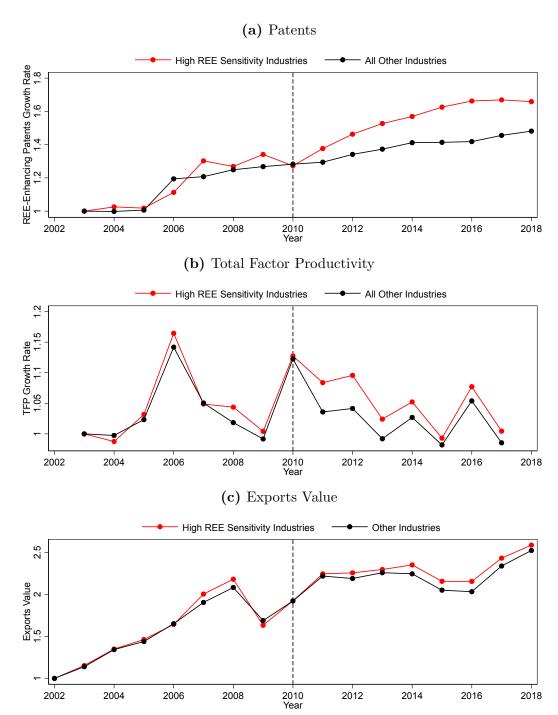
<sup>&</sup>lt;sup>15</sup>In Appendix A.7, we use UN Comtrade trade data disaggregated at the finer 6-digit HS product code level. At this level, we can separate the effect on the growth of export values into price growth and the growth of physical quantities. We find that the positive effect on exports outside China was not driven by price increases but mostly by increases in quantity.

**Table 3:** Downstream Export Growth – Rare-Earth Intense Manufacturing Industries

	Annualized Growth: Exports Value						
	NONCHN	ALL	USA	EUR	JPN	CHN	
	(1)	(2)	(3)	(4)	(5)	(6)	
RE Sens. $\times$ Post	0.880***	0.812***	0.155	0.639*	1.685**	-0.804	
	(0.254)	(0.245)	(0.770)	(0.340)	(0.803)	(0.800)	
Observations	271,740	277,723	6,048	107,895	5,979	5,983	
Clusters	17,249	$17,\!623$	378	6,754	375	374	
	Differences in Annualized Export Growth to China						
	NONCHN	ALL	USA	EUR	JPN		
	(7)	(8)	(9)	(10)	(11)		
RE Sens. $\times$ Post	3.548***	3.382***	3.133***	3.540***	2.920**		
	(0.502)	(0.482)	(0.945)	(0.709)	(1.398)		
Observations	270,342	276,325	5,987	107,321	5,951		
Clusters	17,159	17,533	375	6,722	374		
Controls	Yes	Yes	Yes	Yes	Yes	Yes	
Country x Ind F.E.	Yes	Yes	Yes	Yes	Yes	Yes	
Country x Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes	

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{ist} = \beta$  REE Sens.  $\times$  post $_t + \gamma \Delta_{ist} + \eta_{is} + \eta_{it} + \epsilon_{ist}$ , where  $y_{ist}$  represents the annualized growth rate of export values for country-industry is in year t (upper panel) and the difference between the annualized growth rate of export values for country-industry is and the corresponding growth rate of the same industry s in China (lower panel). The annualized growth is calculated using the midpoint between t and t-1 as the denominator. The sample includes 4-digit SIC manufacturing industries from 2002-2018 across the 50 largest economies. The treatment intensity REE Sens. s is a weighted sum of an REE element-specific complementarity index (ranging from 0 to 100), with weights based on total requirement shares for industry s: REE Sens.  $s = \sum_{e} tr_{es} \times compl_{e}$ . post $_{t}$  is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, Japan, and China. All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and country-industry-specific industrial subsidy fractions from the Global Trade Alert database, all interacted with  $post_t$  and the lagged weighted average growth rate of GDP of the ten largest importers from is. Regressions are weighted by the share of export values  $w_{is} = exports_{is}/\sum_{i,s} exports_{is}$  from 2002-2004. Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.05.

Figure 3: High and Low REE Sensitivity Industries: Patents, Productivity, and Exports



Notes: The figure provides evidence of parallel trends in REE-enhancing patents, TFP, and exports of high REE sensitivity industries and all other industries before 2010. It plots the average, weighted by initial values in 2002-2004, of patents stock growth, TFP growth, and exports values of manufacturing industries that are classified in the top 25 percentile of REE sensitivity (red line), and the same variables for all other manufacturing industries (black line). Growth rates are annualized growth calculated using the midpoint between t and t-1 as the denominator. The sample includes 4-digit SIC manufacturing industries from 2002-2018 across the 50 largest economies excluding China. REE sensitivity is constructed following equation (2). The plotted values are normalized to the value in the base year 2003 (for patents and total factor productivity) or 2002 (for exports value).

Alternative Treatment Measures: While using the index of complementarity in (2) has the appeal that it is directly related to the production functions of REE using industries and the chemical properties of elements, we can use UN Comtrade to construct alternative measures of exposure to the Chinese REE supply shock that vary at the country-industry level instead of varying only across industries. Our first measure of country-industry REE sensitivity uses the country-level price spike in unit values of REE imports interacted with the total requirements of REE of each industry (aggregated across elements):<sup>16</sup>

$$REE \ Sensitivity_{is} = \left(\sum_{e} tr_{es}\right) \times \left(\ln(\max REE \ import \ price_i) - \ln(REE \ import \ price \ 2009_i)\right). \tag{4}$$

Alternatively, we build a second country-industry treatment indicator that uses pre-shock import shares of REE elements in country c sourced from China relative to REEs sourced from a larger set of countries (China, the U.S., Australia, Russia or India) again interacted with the total requirement of REEs of industry s:

$$REE \ Sensitivity_{is} = \left(\sum_{e} tr_{es}\right) \times \frac{REE \ imports \ from \ CHN_{i}}{REE \ imports \ from \ CHN, \ USA, \ AUS, \ RUS, \ IND_{i}}.$$
 (5)

For both shock measures, we consider HS codes 284690, 284610, and 280530 as REE relevant and weigh them by initial import shares. Appendix Table A.8 reports the estimates based on both alternative shock measures using the different regional samples. In line with our previous results, we find that export growth of downstream industries increased relatively more in those country-industries where the REE supply shock was relatively more important.

# 5 A Quantitative Model of Trade and Directed Technological Change

Having presented the empirical evidence, in this section, we develop a model that can generate REE biased innovation and technological change in response to a negative REE supply shock. The model embeds static directed technical change (Acemoglu, 2002) into a quantitative general-equilibrium gravity model of international trade (Caliendo and Parro, 2015; Fadinger et al., 2024). The model features two production factors (tradable REE and non-tradable labor) and flexibly combines standard Heckscher-Ohlin forces of comparative advantage and directed technological change. In response to a REE export tax by China, the REE factor becomes scarcer outside

<sup>&</sup>lt;sup>16</sup>We have used different benchmarks for the price level, obtaining similar results.

of China. Absent directed technological change, this leads to a cost increase that is more pronounced in REE-intensive industries and shifts comparative advantage and production towards labor-intensive industries. By contrast, in the presence of endogenous technology and complementarity between REE and labor inputs, the increase in REE factor costs triggers innovation that is biased towards REEs, leading to productivity gains in REE-intensive sectors relative to labor-intensive ones outside China. This technology effect counteracts Heckscher-Ohlin forces, and, depending on the magnitude of the productivity response, may weaken or dominate the former. In addition, the model features intermediates and inter-industry linkages, which generate amplification of policy changes, as price changes of the REE factor are passed on to downstream customers and other industries. Such amplification is necessary to obtain quantitatively accurate trade and welfare predictions (Caliendo and Parro, 2015).

The production side of the model is structured in four layers: final goods, Armington industry bundles, tradable industry goods, and factor-biased innovation activities. Figure 4 illustrates the model's structure. At the most upstream level, monopolistically competitive innovation firms invent and produce differentiated varieties of labor and REE-specific inputs. These varieties are then used, together with material inputs, to produce country- and industry-specific tradable goods. The tradable goods are aggregated into non-tradable, industry-specific Armington bundles within each country. At the most downstream level, these Armington bundles are further aggregated into non-tradable final goods, which serve either consumption or investment in innovation.

## 5.1 Setup

There are many countries, indexed by i = 1, ..., I and j = 1, ..., J and industries, indexed by s = 1, ..., S. The first country subindex denotes the location of consumption and the second subindex the location of production.

#### 5.1.1 Tradable Industries

Industry goods are country-specific and tradable. The value-added production function of the country-specific industry goods combines quantities of REE and labor input bundles  $Y_{Ris}$  and  $Y_{Lis}$ :

$$VA_{is} = \left[ \gamma_s Y_{Ris}^{\frac{\varepsilon_s - 1}{\varepsilon_s}} + (1 - \gamma_s) Y_{Lis}^{\frac{\varepsilon_s - 1}{\varepsilon_s}} \right]^{\frac{\varepsilon_s}{\varepsilon_s - 1}}$$
(6)

A key parameter in (6) is the industry-specific elasticity of substitution between REE and labor input bundles, denoted by  $\varepsilon_s$ . The value of  $\varepsilon_s$  determines if REE are gross substitutes ( $\varepsilon_s > 1$ ) or

complements ( $\varepsilon_s < 1$ ) with labor in the production of industry value added. This elasticity will later govern the direction of technological change in response to changes in relative factor prices. In addition,  $\gamma_s \in [0,1]$  determines the dependence of industry s on REE input. The higher the value of  $\gamma_s$ , the more dependent a given industry is on the REE bundle.

In a further stage, value added is combined with a Cobb-Douglas aggregate of material bundles  $M_{iss'}$  used by industry s and produced by industries s'. Here,  $\phi_{iss'}$  denotes the IO coefficients (expenditure shares of country i industry s on goods produced by industry s'). The production function for gross output and the resource constraint of the industry goods are given by

$$Y_{is} = \Psi_{is} V A_{is}^{\phi_{is}} \prod_{s'}^{S} M_{iss'}^{\phi_{iss'}} = \sum_{j} d_{jis} Y_{jis}, \tag{7}$$

where  $\Psi_{is} \equiv \phi_{is}^{-\phi_{is}} \prod_{s'} \phi_{iss'}^{-\phi_{iss'}}$  is a constant normalizing the production function. Output  $Y_{is}$  is used in all countries, where  $Y_{jis}$  denotes the quantity of the industry-s good that is produced in country i and used by country j and  $d_{jis} \geq 1$  denotes iceberg-type trade costs with  $d_{iis} = 1$ .

Note that when  $\phi_{is}$  tends to unity, so that material inputs are absent, the model converges to the production structure of Acemoglu (2002). Instead, when  $\gamma_s$  tends to zero, such that labor is the only production factor, we are back to the workhorse quantitative trade model (Caliendo and Parro, 2015).<sup>17</sup>

#### 5.1.2 Innovation and Directed Technological Change

In each country-industry, firms in the  $R_{is}$  and  $L_{is}$  layers are perfectly competitive and use a CES bundle of differentiated inputs  $(y_{Ris}(a) \text{ or } y_{Lis}(a))$  to produce the REE input bundle  $Y_{Ris}$  or the labor input bundle  $Y_{Lis}$ . These inputs cannot be traded across countries, and the technology to produce them does not diffuse. The production functions of the layers  $R_{is}$  and  $L_{is}$  are given by:

$$Y_{Ris} = E_{Ris} \left[ \int_0^{A_{Ris}} y_{Ris}(a)^{\frac{\mu_s - 1}{\mu_s}} da \right]^{\frac{\mu_s}{\mu_s - 1}}, \quad Y_{Lis} = E_{Lis} \left[ \int_0^{A_{Lis}} y_{Lis}(a)^{\frac{\mu_s - 1}{\mu_s}} da \right]^{\frac{\mu_s}{\mu_s - 1}}$$
(8)

Here, the terms  $E_{Ris} = A_{Ris}^{\delta}$  and  $E_{Lis} = A_{Lis}^{\delta}$  are externalities from the measures  $A_{Ris}$ ,  $A_{Lis}$  of input varieties on downstream productivity. If  $\delta > 0$ , spillovers are positive, while if  $\delta < 0$ , there

 $<sup>^{17}</sup>$ We model REE inputs as part of value added rather than as separate material inputs, primarily due to data limitations. Specifically, a highly disaggregated input-output (IO) table that identifies REE as a distinct input is available only for the U.S., but not for other countries. For the purposes of the quantitative analysis, we therefore rely on the WIOD database, which features more aggregated industry classifications and does not report REE-specific IO linkages. This modeling choice allows us to estimate sector-specific elasticities of substitution between REE and labor input bundles,  $\epsilon_s$ , using patent data, and to calibrate the REE expenditure shares,  $\gamma_s$ , based on the U.S. IO table, as described in Section 6.1.

are negative spillovers. If  $\delta = 0$ , spillover effects are absent.<sup>18</sup>

The corresponding price indices for the REE and labor input bundles are:

$$P_{Ris} = E_{Ris}^{-1} \left[ \int_0^{A_{Ris}} p_{Ris}(a)^{1-\mu_s} da \right]^{\frac{1}{1-\mu_s}}, \quad P_{Lis} = E_{Lis}^{-1} \left[ \int_0^{A_{Lis}} p_{Lis}(a)^{1-\mu_s} da \right]^{\frac{1}{1-\mu_s}}$$
(9)

The inverse demand faced by an input firm that produces a given input variety a is:

$$p_{Ris}(a) = E_{Ris}^{\frac{(\mu_s - 1)}{\mu_s}} P_{Ris} Y_{Ris}^{\frac{1}{\mu_s}} y_{Ris}(a)^{\frac{-1}{\mu_s}}, \quad p_{Lis}(a) = E_{Ris}^{\frac{(\mu_s - 1)}{\mu_s}} P_{Lis} Y_{Lis}^{\frac{1}{\mu_s}} y_{Lis}(a)^{\frac{-1}{\mu_s}}$$
(10)

The measures  $A_{Ris}$  of REE input firms and  $A_{Lis}$  of labor input firms denote the state of technology. They are endogenous and determined by free entry into innovation. Input firms only sell domestically, operate under monopolistic competition, and hold a patent for their variety. For simplicity, we assume that input firms and their patent die after one period and are replaced by new entrants. Varieties of inputs are imperfect substitutes with elasticity of substitution  $\mu_s > 1$ . Each REE (labor) input variety is produced with a linear technology with the factor REE (labor)  $r_{is}$  ( $l_{is}$ ) as input:

$$y_{Ris}(a) = r_{is}(a), \quad y_{Lis}(a) = l_{is}(a)$$
 (11)

Input firms maximize profits, taking their inverse demand (10) and production technology (11) as given. Solving their profit-maximization problem yields the optimal prices of inputs

$$p_{Ris}(a) = p_{Ris} = \frac{\mu_s}{\mu_s - 1} w_{Ri}, \quad p_{Lis}(a) = p_{Lis} = \frac{\mu_s}{\mu_s - 1} w_{Li},$$
 (12)

which correspond to a markup over factor prices  $w_{Ri}$  ( $w_{Li}$ ), and and the variable profits of input monopolists

$$\pi_{Ris} = \frac{p_{Ris}r_{is}}{\mu_s} = \frac{1}{\mu_s} \left(\frac{\mu_s}{\mu_s - 1}\right)^{1-\mu_s} P_{Ris}^{\mu_s} E_{Ris}^{\mu_s - 1} Y_{Ris} w_{Ri}^{1-\mu_s},$$

$$\pi_{Lis} = \frac{p_{Lis}l_{is}}{\mu_s} = \frac{1}{\mu_s} \left(\frac{\mu_s}{\mu_s - 1}\right)^{1-\mu_s} P_{Lis}^{\mu_s} E_{Ris}^{\mu_s - 1} Y_{Lis} w_{Li}^{1-\mu_s}.$$
(13)

Inventing an input patent is associated with fixed costs  $f_{Ri}$  (or  $f_{Li}$ ) which are paid in units of the

<sup>&</sup>lt;sup>18</sup>The formulation of these spillover effects follows Benassy (1996) and allows disentangling love for variety from markups. Spillover effects are absent in Acemoglu (2002) and not required to generate directed technological change in response to changes in factor prices. However, they determine the strength of the technology response and the response of revenue in downstream industries to changes in relative factor prices. In particular, negative spillovers increase the technology response to changes in factor prices by increasing prices of input bundles and thus profits of innovators if inputs are complements (see equation (43) below). For example, they could result from (unmodeled) competition for a scarce factor.

final good in country i with price  $P_i$ . Note that since  $P_{Ris} = E_{Ris}^{-1} A_{Ris}^{\frac{1}{1-\mu_s}} p_{is}$  and  $Y_{Ris} = E_{Ris} A_{Ris}^{\frac{\mu_s}{\mu_s-1}} r_{is}$ , we have that  $p_{Ris}r_{is} = P_{Ris}Y_{Ris}/A_{Ris}$  and similarly,  $p_{Lis}l_{is} = P_{Lis}Y_{Lis}/A_{Lis}$ . Free entry implies that inventing a new patent allows input firms to exactly recoup the innovation fixed cost:

$$\Pi_{Ris} = \frac{p_{Ris}r_{is}}{\mu_s} = \frac{P_{Ris}Y_{Ris}}{A_{Ris}\mu_s} = f_{Ris}P_i, \quad \Pi_{Lis} = \frac{p_{Lis}l_{is}}{\mu_s} = \frac{P_{Lis}Y_{Lis}}{A_{Lis}\mu_s} = f_{Lis}P_i$$
 (14)

Taking ratios of the free-entry conditions yields the relative technology bias:

$$\frac{A_{Ris}}{A_{Lis}} = \frac{P_{Ris}Y_{Ris}}{P_{Lis}Y_{Lis}} \frac{f_{Lis}}{f_{Ris}} \tag{15}$$

Thus, the relative technology bias depends on the relative revenues of firms in the  $R_{is}$  compared to the  $L_{is}$  sector and any change in relative revenues will shift the relative technology bias.

#### 5.1.3 Final Goods and Armington Industry Bundles

We now specify the remainder of the model. At the most downstream level, there is a non-tradable final good  $Y_i$ , that is produced and used in each country. This final good can be used for two purposes. A part of its production serves final consumer demand (denoted by  $C_i$ ), while the remainder is used to pay the fixed costs of innovation (denoted by  $\sum_s A_{Ris} f_{Ris}$  and  $\sum_s A_{Lis} f_{Lis}$ ). The final good is produced with a CES production function that combines Armington bundles of industry goods. The production function and resource constraint of the final good are given by

$$Y_{i} = \left[\sum_{s} \alpha_{is}^{\frac{1}{\rho}} Q_{is}^{\frac{\rho-1}{\rho}}\right]^{\frac{\rho}{\rho-1}} = C_{i} + \sum_{s} A_{Ris} f_{Ris} + \sum_{s} A_{Lis} f_{Lis},$$
(16)

where  $Q_{is}$  corresponds to the quantity of industry-s Armington bundle used in the production of the final good.

The Armington bundles  $M_{is}$  are produced with a CES technology that aggregates the tradable industry-s inputs  $Y_{ijs}$  from all countries j = 1, ..., J with an elasticity of substitution  $\sigma_s > 1$ . These Armington bundles serve two purposes. Either they are used as an input in the production of final goods (as  $Q_{is}$ ), or they are used as intermediates in the production of tradable goods (as  $M_{iss'}$ ). The production function and resource constraint of the Armington goods are given by

$$M_{is} = \left[\sum_{j} Y_{ijs}^{\frac{\sigma_s - 1}{\sigma_s}}\right]^{\frac{\sigma_s}{\sigma_s - 1}} = Q_{is} + \sum_{s'} M_{iss'}, \tag{17}$$

where  $M_{iss'}$  is the quantity of industry-s Armington bundles used in industry s' within country i.

The associated CES price index of the industry-s Armington bundle in country i is

$$P_{is} = \left[\sum_{j} P_{ijs}^{1-\sigma_s}\right]^{\frac{1}{1-\sigma_s}},\tag{18}$$

with  $P_{ijs}$  being the price of the country-specific industry s goods used in country i and produced in country j. The aggregate price index of the final good in country i is

$$P_i = \left[\sum_s \alpha_{is} P_{is}^{1-\rho}\right]^{\frac{1}{1-\rho}}.$$
 (19)

Assuming perfect competition in Armington bundles, the bilateral trade value shares, measuring the import values of i from j relative to the total expenditure in country i, are given by:

$$\lambda_{ijs} = \frac{P_{ijs}Y_{ijs}}{P_{is}M_{is}} = \left(\frac{P_{ijs}}{P_{is}}\right)^{1-\sigma_s}.$$
 (20)

#### 5.1.4 Trade and Industrial Policy

We assume that China is the only country that has endowments of a homogeneous REE factor  $R_C$  that can be traded. It additionally disposes of a non-discriminatory gross ad-valorem export tax  $\tau_{XC} \geq 1$  on the exports of REEs. Therefore, the price of REE inputs in country j is equal to  $w_{Rj} = \tau_{XC} w_{RC}$ .

#### 5.1.5 Goods and Factor Markets

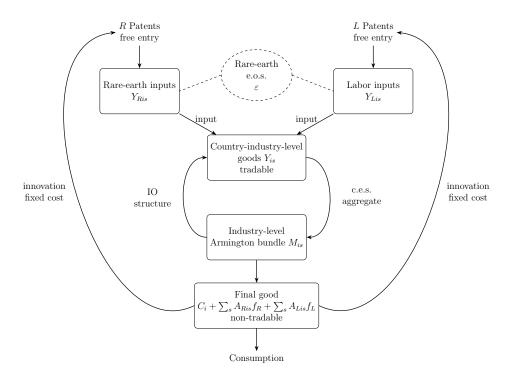
Denote the revenues of industry s in country i by  $Rev_{is} = P_{iis}Y_{is}$ . Product markets for industry Armington bundles clear if the industry's revenue equals expenditure on the industry's goods:

$$Rev_{is} = \sum_{j} \lambda_{jis} \left[ \alpha_{js} \left( \frac{P_{js}}{P_{j}} \right)^{1-\rho} \left( w_{Lj} L_{j} + I_{C} \times w_{RC} R_{C} + T_{j} - T B_{j} - NFI A_{j} \right) + \sum_{s'} \phi_{ss'j} Rev_{s'j} \right]. \tag{21}$$

Here,  $I_C$  is an indicator variable that equals unity when j is China and zero for all other countries j.  $TB_j$  is the (exogenous) trade imbalance and

$$NFIA_{j} = \begin{cases} -w_{Rj} \sum_{s=1}^{S} \int_{0}^{A_{Rjs}} r_{js}(a) da, & \text{if } j \neq C \\ \sum_{j \neq C} w_{Rj} \sum_{s=1}^{S} \int_{0}^{A_{Rjs}} r_{js}(a) da, & \text{if } j = C \end{cases}$$
(22)

Figure 4: Model Structure



*Notes:* The figure depicts the general equilibrium structure of the model.

are net factor incomes from abroad,  $NFIA_j$ , arising from the use of the tradable REE factor for countries other than China and for China, respectively. Total export tax revenue on net exports of REEs are fully rebated to consumers with lump-sum transfers  $T_i$ :

$$T_j = I_C \times (\tau_{XC} - 1) w_{RC} (R_C - \sum_s A_{RjC} r_{Cs})$$

Labor is immobile across countries, and labor markets clear for each country:

$$\sum_{s} \int_{0}^{A_{Lsi}} l_s(a) da = L_i \tag{23}$$

By contrast, the REE factor is tradable and the REE market clears at the world level:

$$\sum_{j} \sum_{s} \int_{0}^{A_{Rjs}} r_{js}(a)da = R_C \tag{24}$$

## 5.2 Equilibrium

An equilibrium in this model determines a solution for (i) factor prices  $w_{RC}$  for China and  $w_{Li} \forall i$ ; (ii) revenues, prices and trade shares for tradable country-industry pairs  $Rev_{is} \forall i, s, P_{is} \forall i, s, P_{i} \forall i, s, P_{i} \forall i, s, P_{i} \forall i, j, s, P_{i} \forall i, j,$ 

This solution is determined by (i)  $I \times S$  equations (25), determining  $Rev_{is}$ ; (ii)  $I \times J \times S$  equations (28) determining  $\lambda_{ijs}$ ; (iii)  $I \times J \times S$  equations (29) determining  $P_{ijs}$ ; (iv)  $I \times S$  equations (30) determining  $P_{is}$ ; (v) I equations (31) determining  $P_i$ ; (vi)  $I \times I \times S$  equations of type (32) determining  $P_{Ris}Y_{Ris}$  and  $P_{Lis}Y_{Lis}$ ; (vii) I equations (35) determining  $w_{Li}$  and (viii) 1 equation (36), determining  $w_{RC}$ . Further, note that  $w_{Ri} = w_{RC}\tau_{XC}$ .

Market clearing for Armington goods is given by:

$$Rev_{is} = \sum_{j} \lambda_{jis} \times$$

$$\left[ \alpha_{js} \left( \frac{P_{js}}{P_{j}} \right)^{1-\rho} \left( w_{Lj} L_{j} + I_{C} \times w_{RC} R_{C} + T_{j} - T B_{j} - N F I A_{j} \right) + \sum_{s} \phi_{ss'j} Rev_{js} \right] \quad \forall i, s,$$

$$(25)$$

with net factor incomes from abroad being:

$$NFIA_{j} = \begin{cases} -\sum_{s=1}^{S} \left(\frac{\mu_{s}-1}{\mu_{s}}\right) \left(P_{Rjs}Y_{Rjs}\right), & \forall j \neq C \\ \sum_{j\neq C} \sum_{s=1}^{S} \left(\frac{\mu_{s}-1}{\mu_{s}}\right) \left(P_{Rjs}Y_{Rjs}\right), & \forall j = C \end{cases}$$

$$(26)$$

Transfers  $T_j$  are given by:

$$T_j = I_C \times (\tau_{XC} - 1)[w_{RC}R_C - \sum_s \left(\frac{\mu_s - 1}{\mu_s}\right) \tau_{RCs}^{-1}(P_{RCs}Y_{RCs})], \tag{27}$$

where  $I_C$  is an indicator that equals unity for j = China.

Bilateral trade shares are:

$$\lambda_{ijs} = \left(\frac{P_{ijs}}{P_{is}}\right)^{1-\sigma_s} \quad \forall i, j, s \tag{28}$$

Bilateral prices are given by:

$$P_{ijs} = d_{ijs} \left[ \gamma_s^{\varepsilon_s} \left( \frac{P_{Rjs} Y_{Rjs}}{\mu_s f_{Rjs} P_j} \right)^{\delta(\varepsilon_s - 1) + \frac{\varepsilon_s - 1}{\mu_s - 1}} \left( \frac{\mu_s}{\mu_s - 1} \right)^{1 - \varepsilon_s} (w_{Rj})^{1 - \varepsilon_s} +$$

$$(1 - \gamma_s)^{\varepsilon_s} \left( \frac{P_{Ljs} Y_{Ljs}}{\mu_s f_{Ljs} P_j} \right)^{\delta(\varepsilon_s - 1) + \frac{\varepsilon_s - 1}{\mu_s - 1}} \left( \frac{\mu_s}{\mu_s - 1} \right)^{1 - \varepsilon_s} (w_{Lj})^{1 - \varepsilon_s} \right]^{\frac{\phi_{is}}{1 - \varepsilon_s}} \prod_{s'} P_{js}^{\phi_{jss'}} \quad \forall i, j, s$$

Industry-level price indices are defined as:

$$P_{is} = \left[\sum_{j} P_{ijs}^{1-\sigma_s}\right]^{\frac{1}{1-\sigma_s}} \quad \forall i, s$$
(30)

Aggregate price levels are given by:

$$P_i = \left[\sum_s \alpha_{is} P_{is}^{1-\rho}\right]^{\frac{1}{1-\rho}} \quad \forall i \tag{31}$$

Revenues of the  $R_{is}$  and  $L_{is}$  layers are given by:

$$P_{Ris}Y_{Ris} = \gamma_s \phi_{is}Rev_{is} \times$$

$$\left[ \gamma_s + (1 - \gamma_s) \left( \frac{(1 - \gamma_s)f_{Ris}}{\gamma_s f_{Lis}} \right)^{\frac{(\varepsilon_s - 1)\{\mu_s \kappa_s + (\mu_s - 1)^2 \varepsilon_s \delta\}}{(\mu_s - \varepsilon_s)[\mu_s - \varepsilon_s - \delta(\mu_s - 1)(\varepsilon_s - 1)]}} \left( \frac{r_{is}}{l_{is}} \right)^{\frac{(1 - \varepsilon_s)\{\delta(\mu_s - 1)^2(\varepsilon_s - 1) + (\mu_s - 1)\kappa_s\}}{(\mu_s - \varepsilon_s)\kappa_s}} \right]^{-1} \forall i, s$$

$$P_{Lis}Y_{Lis} = (1 - \gamma_s)\phi_{is}Rev_{is} \times$$

$$\left[ (1 - \gamma_s) + \gamma_s \left( \frac{(1 - \gamma_s)f_{Ris}}{\gamma_s f_{Lis}} \right)^{\frac{(1 - \varepsilon_s)\{\mu_s \kappa_s + (\mu_s - 1)^2 \varepsilon_s \delta\}}{(\mu_s - \varepsilon_s)\kappa_s}} \left( \frac{r_{is}}{l_{is}} \right)^{\frac{(\varepsilon_s - 1)\{\delta(\mu_s - 1)^2(\varepsilon_s - 1) + (\mu_s - 1)\kappa_s\}}{(\mu_s - \varepsilon_s)\kappa_s}} \right]^{-1} \forall i, s,$$

where  $\kappa_s \equiv \mu_s - \varepsilon_s + \delta(\mu_s - 1)(1 - \varepsilon_s)$  and

$$\frac{r_{is}}{l_{is}} = \frac{f_{Ris}}{f_{Lis}} \frac{w_{Li}}{w_{Ri}} \quad \forall i, s. \tag{34}$$

Labor markets clear in each country:

$$\sum_{s} \left( \frac{\mu_s - 1}{\mu_s} \right) (P_{Lis} Y_{Lis}) = w_{Li} L_i \quad \forall i$$
 (35)

The REE market clears at the world level:

$$\sum_{j} \sum_{s} \left( \frac{\mu_{s} - 1}{\mu_{s}} \right) (P_{Rjs} Y_{Rjs}) \tau_{XC}^{-1} = w_{RC} R_{C}$$
(36)

The Model with Exogenous Technology: For comparison, we also consider a version of the model where  $A_{Ris}$  and  $A_{Lis}$  are exogenously given and do not respond to policy. In this case, the free entry conditions do not hold. With exogenous technology, equilibrium equations (32) need to

be replaced by:

$$P_{Ris}Y_{Ris} = \gamma_s \phi_{is} Rev_{is} \times$$

$$\left[ \gamma_s + (1 - \gamma_s) \left( \frac{\gamma_s}{1 - \gamma_s} \right)^{1 - \varepsilon_s} \left( \frac{A_{Ris}}{A_{Lis}} \right)^{(1 - \varepsilon_s) \left( \delta + \frac{\mu_s}{\mu_s - 1} - 1 \right)} \left( \frac{w_{Li}}{w_{Ri}} \right)^{1 - \varepsilon_s} \right]^{-1}$$

$$P_{Lis}Y_{Lis} = (1 - \gamma_s) \phi_{is} Rev_{is} \times$$

$$\left[ (1 - \gamma_s) + \gamma_s \left( \frac{\gamma_s}{1 - \gamma_s} \right)^{\varepsilon_s - 1} \left( \frac{A_{Ris}}{A_{Lis}} \right)^{(\varepsilon_s - 1) \left( \delta + \frac{\mu_s}{\mu_s - 1} - 1 \right)} \left( \frac{w_{Li}}{w_{Ri}} \right)^{\varepsilon_s - 1} \right]^{-1}$$

$$(38)$$

Moreover, (29) becomes:

$$P_{ijs} = d_{ijs} \left[ \gamma_s^{\varepsilon_s} \left( A_{Ris} \right)^{\delta(\varepsilon_s - 1) + \frac{\varepsilon_s - 1}{\mu_s - 1}} \left( \frac{\mu_s}{\mu_s - 1} \right)^{1 - \varepsilon_s} w_{Rj}^{1 - \varepsilon_s} + \left( 1 - \gamma_s \right)^{\varepsilon_s} \left( A_{Lis} \right)^{\delta(\varepsilon_s - 1) + \frac{\varepsilon_s - 1}{\mu_s - 1}} \left( \frac{\mu_s}{\mu_s - 1} \right)^{1 - \varepsilon_s} w_{Lj}^{1 - \varepsilon_s} \right]^{\frac{\phi_{is}}{1 - \varepsilon_s}} \prod_{s'} P_{js}^{\phi_{jss'}}$$

$$(39)$$

The remaining equilibrium conditions are unaffected.

#### 5.3 Discussion

**Direction of Innovation Bias:** Using the model, we can analytically determine the direction in which the innovation bias, measured by the ratio  $\frac{A_{Ris}}{A_{Lis}}$ , shifts in response to changes in the factor input ratio  $\frac{r_{is}}{l_{is}}$  for each country-industry pair is.

To illustrate this, we first take the ratio of the production functions for the  $R_{is}$  and  $L_{is}$  input layers, as given in equation (8):

$$\frac{Y_{Ris}}{Y_{Lis}} = \frac{E_{Ris}}{E_{Lis}} \left(\frac{A_{Ris}}{A_{Lis}}\right)^{\frac{\mu_s}{\mu_s - 1}} \frac{r_{is}}{l_{is}}.$$
(40)

Next, combining this expression with the relative demand equation

$$\frac{Y_{Ris}}{Y_{Lis}} = \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\varepsilon} \left(\frac{P_{Ris}}{P_{Lis}}\right)^{-\varepsilon},\tag{41}$$

we derive an expression for the relative revenues of firms in the  $R_{is}$  and  $L_{is}$  sectors:

$$\frac{P_{Ris}Y_{Ris}}{P_{Lis}Y_{Lis}} = \frac{\gamma_s}{1 - \gamma_s} \left(\frac{E_{Ris}}{E_{Lis}}\right)^{\frac{\varepsilon_s - 1}{\varepsilon_s}} \left(\frac{A_{Ris}}{A_{Lis}}\right)^{\frac{\mu_s}{\mu_s - 1}} \frac{\varepsilon_s - 1}{\varepsilon_s} \left(\frac{r_{is}}{l_{is}}\right)^{\frac{\varepsilon_s - 1}{\varepsilon_s}}$$
(42)

Combining (42) with (15), and assuming that  $E_{Ris}/E_{Lis} = (A_{Ris}/A_{Lis})^{\delta}$ , we can solve for the relative technology bias as a function of the input ratio:

$$\frac{A_{Ris}}{A_{Lis}} = \left(\frac{f_{Lis}}{f_{Ris}} \frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{(\mu_s - 1)\varepsilon_s}{\kappa_s}} \left(\frac{r_{is}}{l_{is}}\right)^{\frac{(\mu_s - 1)(\varepsilon_s - 1)}{\kappa_s}}$$
(43)

When  $\mu_s > 1$  (a necessary condition to ensure well-defined profit maximization by monopolistic producers) and  $\varepsilon_s < 1$  (indicating that REEs and labor are gross complements), the expression in (43) is decreasing in  $\frac{r_{is}}{l_{is}}$  if  $\delta \geq 0$ , or if  $|\delta| < \frac{\mu_s - \varepsilon_s}{(\mu_s - 1)(\varepsilon_s - 1)}$  and  $\delta < 0$ . Under these conditions, a negative REE supply shock — i.e., a decline in  $\frac{r_{is}}{l_{is}}$  — biases innovation toward increasing the efficiency of REE inputs.

The underlying intuition is as follows: with complementary inputs, the price effect (arising from an increase in  $P_{Ris}/P_{Lis}$ ) dominates the negative market size effect (due to reduced  $r_{is}/l_{is}$ ). Consequently, developing technologies that enhance the productivity of the REE-intensive sector  $R_{is}$  becomes relatively more profitable than innovations in the labor-intensive sector  $L_{is}$ .

Moreover, provided that  $\kappa_s > 0$ , the sensitivity of the technology bias to changes in  $\frac{r_{is}}{l_{is}}$  increases with the magnitude of negative  $\delta$ . This reflects the role of negative spillovers: when inputs are complementary, such spillovers reduce the productivity of competitors, raising the price of the input bundle  $(P_{Ris}/P_{Lis})$  and thereby increasing the profitability of innovation despite falling output levels. This effect is captured in the expression for relative revenue (44), which results from substituting (43) into (42):

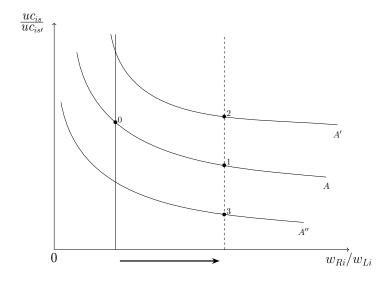
$$\frac{P_{Ris}Y_{Ris}}{P_{Lis}Y_{Lis}} = \left(\frac{f_{Lis}}{f_{Ris}}\right)^{\frac{\mu_s(\varepsilon_s - 1)\kappa_s + (\mu_s - 1)^2\varepsilon_s\delta(\varepsilon_s - 1)}{(\mu_s - \varepsilon_s)\kappa_s}} \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{\varepsilon_s(\mu_s - 1)\kappa_s + (\mu_s - 1)^2\varepsilon_s\delta(\varepsilon_s - 1)}{(\mu_s - \varepsilon_s)\kappa_s}} \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{\varepsilon_s(\mu_s - 1)\kappa_s + (\mu_s - 1)^2\varepsilon_s\delta(\varepsilon_s - 1)}{(\mu_s - \varepsilon_s)\kappa_s}} \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{\varepsilon_s(\mu_s - 1)\kappa_s + (\mu_s - 1)^2\varepsilon_s\delta(\varepsilon_s - 1)}{(\mu_s - \varepsilon_s)\kappa_s}} \right) \tag{44}$$

Similar to the case of technology bias, the expression for relative revenue is decreasing in  $\frac{r_{is}}{l_{is}}$  under analogous conditions, implying that a negative REE supply shock raises the relative profitability of the  $R_{is}$  sector.

By contrast, the relative output of the REE sector,  $Y_{Ris}/Y_{Lis}$ , increases with the input ratio  $\frac{r_{is}}{l_{is}}$ :

$$\frac{Y_{Ris}}{Y_{Lis}} = \left(\frac{f_{Lis}}{f_{Ris}} \frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{\mu_s \varepsilon_s \kappa_s + \delta(\mu_s - 1)^2 \varepsilon_s^2}{(\mu_s - \varepsilon_s)\kappa_s}} \left(\frac{r_{is}}{l_{is}}\right)^{\frac{\delta(\mu_s - 1)^2 \varepsilon_s (\varepsilon_s - 1) + (\mu_s - 1)\varepsilon_s \kappa_s}{(\mu_s - \varepsilon_s)\kappa_s}}$$
(45)

Figure 5: Factor Prices and the Direction of Technological Change



Notes: The figure plots possible adjustments in response to an increase in the relative REE factor price for two industries s and s', where industry s' is more reliant on REE. Curves A, A' and A'' depict different relative unit costs depending on the direction of technological change.

Finally, the optimal factor input ratio is inversely related to the relative price of REEs:

$$\frac{r_{is}}{l_{is}} = \frac{f_{Ris}}{f_{Lis}} \frac{w_{Li}}{w_{Ri}}$$

Impact on Comparative Advantage: By comparing the expressions for unit costs under endogenous (29) and exogenous technologies (39), we gain insight into how the competitiveness of downstream industries and comparative advantage responds to an increase in the price of the REE factor.

Under exogenous technology, an increase in  $w_{Rj}$  raises production costs — an effect that is more pronounced in industries with higher REE intensity (factor cost effect). According to the trade-share equation (20), this reduces bilateral exports of REE-intensive industries relative to labor-intensive ones. This effect captures standard Heckscher-Ohlin forces of comparative advantage.

By contrast, under endogenous technology, an increase in  $w_{Rj}$  also induces a technological response via an associated increase in  $P_{Ris}Y_{Ris}$  (the technology effect). If this effect is sufficiently strong, it can outweigh the factor cost effect, thereby enhancing the relative competitiveness of REE-intensive industries. As a result, exports of these industries may increase relative to those of labor-intensive industries despite the rise in input costs.

This can be illustrated with the help of Figure 5. The figure shows relative unit costs for two

industries s and s' and highlights potential equilibrium adjustments in response to an increase in the relative price of REE. Since these relative unit-cost curves are downward sloping in the relative price of REE  $w_{Ri}/w_{Li}$ , sector s' is relatively more REE-intensive. Starting from an initial equilibrium at point 0, an increase in the price of REE leads to a new equilibrium with new relative unit costs. When technologies are exogenously fixed, the adjustment occurs along the A curve to the new equilibrium point 1: the REE-intensive sector loses some of its competitiveness compared to the labor-intensive one due to the relatively intensive use of the REE factor, whose relative price has risen.

However, if the response of the technology bias towards REE is sufficiently strong, the new relative unit-cost curve shifts upwards (denoted as A') such that at this alternative equilibrium point 2, the competitiveness of the REE-intensive industry increases relative to the labor-intensive one. Alternatively, if the bias of innovation were to shift in the opposite direction (e.g. if REE and labor were gross substitutes), relative unit costs would become A''. In this alternative equilibrium (point 3), the technology response would amplify standard Heckscher-Ohlin forces and the REE-intensive industry would further lose relative competitiveness compared to the case of exogenous technology. As our model flexibly nests all these cases of the response of comparative advantage to an export tax on REEs, we need to calibrate the model in order to obtain the qualitative and quantitative general-equilibrium response of trade, production and welfare.

# 6 Quantification

In this section, we first calibrate the model to a baseline economy that matches moments from before the REE export restrictions. We then provide a quantitative assessment of the REE crisis and study the impact that the REE crisis would have had in a counterfactual economy without endogenous technological change.

# 6.1 Taking the Model to the Data

We calibrate the model to a baseline economy in 2009, the year before the REE supply shock. We aggregate the world IO tables to 5 regions: China, Europe, Japan, the U.S., and the rest of the world (RoW), including all remaining countries. We consider 14 industries: 12 sectors in manufacturing, agriculture and an aggregated service sector.

Estimating the Substitution Elasticity for REE: A key object in our calibration is the industry-specific substitution elasticity between REE and labor ( $\varepsilon_s$ ) since this elasticity governs

the response of input costs and innovation to a REE supply shock. To shift the direction of innovation towards REEs in response to a positive REE price shock, like in the empirical findings, REEs need to be gross complementary inputs ( $\varepsilon_s < 1$ ). While input elasticities of substitution are typically estimated from variation in input expenditure shares this information is not available at the country-industry level for REEs. Instead, we exploit the model structure to identify this elasticity of substitution from the relation between relative factor prices and the shift in the direction of innovation in each industry.

Our model implies the following structural relationship between relative patents  $A_{Ris}/A_{Lis}$  and relative factor prices  $w_{Ri}/w_{Li}$  that we exploit to estimate  $\varepsilon_s$ :

$$\log\left(\frac{A_{Ris}}{A_{Lis}}\right) = \beta_s \log\left(\frac{w_{Ri}}{w_{Li}}\right) + \delta_s + u_{is},\tag{46}$$

where the coefficient of interest is  $\beta_s \equiv \frac{(1-\varepsilon_s)(\mu_s-1)}{\kappa_s}$ ,  $\delta_s \equiv \frac{\varepsilon_s(\mu_s-1)}{\kappa_s}\log\left(\frac{\gamma_s}{1-\gamma_s}\right)$  is an industry fixed effect, and the structural error term is  $u_{is} \equiv \frac{\mu_s-1}{\kappa_s}\log\left(\frac{f_{Lis}}{f_{Ris}}\right)$ . This equation can be estimated from data on the relative number of REE patents, REE unit import values and wages. We use our patent data, classified with the LLM algorithm described in Section 3.2. The REE unit value for each country is obtained by dividing the value of REE imports by that country with the physical quantity. Finally, labor costs for each country are calculated from the Penn World Tables. We estimate (46) separately for each 2-digit manufacturing SIC code, using a panel of our 5 regions over the sample period from 2002 to 2018. We then aggregate the resulting REE elasticity of substitution estimates to the level of WIOD industries by taking averages weighted by value added.

Calibrating the REE Intensity: Another important parameter in the quantification exercise is the REE intensity of each industry  $\gamma_s$ . From cost minimization, we obtain the following expression for the REE expenditure share in value added:

$$\frac{P_{Ris}Y_{Ris}}{P_{VAis}VA_{is}} = \frac{1}{1 + \left(\frac{1-\gamma_s}{\gamma_s}\right)^{1/\varepsilon_s} \left(\frac{P_{Lis}}{P_{Ris}}\right)^{1-\varepsilon_s}}.$$
(47)

Given the sectoral total requirements of REE for the U.S. from our IO table presented in Section 3.1, combined with data on value added, prices  $P_{Ris}$  and  $P_{Lis}$ , and estimates for  $\varepsilon_s$ , we can back out  $\gamma_s$  for all industries s.<sup>20</sup>

<sup>&</sup>lt;sup>19</sup>We refer to Appendix B.1 for a derivation.

<sup>&</sup>lt;sup>20</sup>REE total requirements represent the value of REE required for each dollar of *final* demand. Therefore, we can obtain the ratio of REE expenditure to value added on the left-hand side by multiplying the REE total requirement

Table 4 presents the estimates of  $\varepsilon_s$  and  $\gamma_s$  by industry, sorted by descending values of  $\varepsilon_s$ . We estimate  $\varepsilon_s \in [0.674, 1.215]$ , with point estimates below unity in most industries. It stands out that those industries with the lowest substitution elasticity are also the most REE-intensive industries. Given these estimates for  $\gamma_s$  and  $\varepsilon_s$ , we iterate in the calibration over the innovation fix costs to exactly match the relative patents  $\frac{A_{Ris}}{A_{Lis}}$  for each country-industry pair in the baseline economy.

Matching World Input-Output Data: For each country-industry pair, we match bilateral trade shares  $\lambda_{ijs}$ , consumption expenditure shares  $\alpha_{is} (P_{is}/P_i)^{1-\rho}$ , the value-added shares in gross output  $\phi_{is}$  and the direct requirement shares in gross output  $\phi_{iss}$ , directly from the data, using trade and production data from the World Input-Output Tables (WIOD, 2012 release). At the country level, we calibrate the size of labor endowments across countries according to country-level employment data adjusted for human-capital differences from the Penn World Tables. We calibrate the model to fit real GDP differences across countries, taking differences in price levels  $P_i$  across economies into account which we also obtain from the Penn World Tables.

Aggregate Parameters: We set the trade elasticity  $\sigma_s$  equal to 6 for all industries, a standard value in the literature (Costinot and Rodríguez-Clare, 2014). We choose a substitution elasticity in the innovation layer  $\mu_s$  of 6.5, implying a markup of 18%, consistent with estimates from De Loecker and Eeckhout (2018). The elasticity of final demand is set to 1.36, following Redding and Weinstein (2024). We calibrate the parameter  $\delta$  that governs the strength of innovation spillovers to -0.1 to match the empirical estimate of the policy on directed technological change across sectors outside of China.<sup>21</sup>

# 6.2 Quantitative Assessment of the Rare-Earth Crisis

Export Tax and Supply Restriction: Starting from the calibrated baseline world economy, we are now in the position to provide a quantitative assessment of the Chinese industrial policy regarding REEs. Consistent with the fact that China introduced export restrictions in combination with an overall reduction in REE supply, we consider the introduction of an REE export tax on all countries, combined with a reduction of REE supply. Since the direct parameters of the Chinese REE industrial policy are not observable, we need to identify them from observable data. While

for each industry with its ratio of final demand to value added, calculated from the supply-use table. Since this ratio is based on U.S. data, we also use U.S. data for labor cost, REE unit import prices and patent data to calculate the ideal price indices  $P_{Lis}$  and  $P_{Ris}$ .

<sup>&</sup>lt;sup>21</sup>We regress the change in relative patents in the calibrated data on sectoral REE sensitivity, using normalized values of  $\frac{\gamma_s}{\varepsilon_s}$  and compare this estimate with a normalized estimate from column (1) of Table 1 as the empirical counterpart.

**Table 4:** Estimates of  $\varepsilon_s$  and  $\gamma_s$ 

Manufacturing Industry	$\varepsilon_s$	$\gamma_s$
Transport equipment	0.674	0.00278
Basic metals and fabricated metal	0.755	0.00319
Mining, petroleum and coal products	0.837	0.00112
Rubber and plastics	0.839	0.00021
Chemicals and chemical products	0.939	0.00003
Other non-metallic mineral products	0.959	0.00009
Machinery	0.976	0.00052
Computer and electronic products	0.985	0.00028
Wood and paper products	1.072	0.00010
Food, beverages and tobacco	1.138	0.00004
Furniture and misc. manufacturing	1.205	0.00003
Textiles and textile products	1.215	0.00003

Notes: The table shows estimates of the REE elasticity of substitution  $\varepsilon_s$  and REE intensity  $\gamma_s$  used in the quantification.

mine output contracted by about 25% (see Figure A.1 in the Appendix), additional tightening arose from reduced refined REE compounds as well as export quotas for certain REE-related HS codes. Likewise, the mix of Chinese trade policy instruments does not map into a single statutory export tax rate. We therefore discipline the policy using two statistics on prices. First, we infer an export tax equivalent of  $\tau_{XC} = 8$  from the average wedge between FOB and EXW REE prices (measured in RMB), which captures export-related frictions. Second, note that the export tax  $\tau_{XC}$  only determines the REE price wedge between China and the rest of the world but not the global level of scarcity. We therefore choose the REE supply restriction for a given  $\tau_{XC} = 8$  to match the average increase of relative factor prices  $\frac{w_{Ri}}{w_{Li}}$  outside of China, observed over a longer horizon between the pre-recession year 2007 and 2017. This yields a 40% reduction in global REE supply  $R_C$ .

Directed Technological Change and Exports: Figure 6 visualizes the innovation response and subsequent change in exports across industries in response to the change in Chinese REE policies. Panel (a) depicts directed technological change, measured by the relative change in the industry-level innovation bias  $\Delta(A_{Ris}/A_{Lis})$  and averaged across countries outside of China. There is substantial heterogeneity in the shift of the innovation bias. As discussed before in Subsection 5.3, the shift in the innovation bias towards the REE factor is stronger, the lower the elasticity of substitution  $\varepsilon_s$  of an industry is. In industries where REEs are a complementary input ( $\varepsilon_s < 1$ ), the innovation bias shifts towards REE. In the other industries, the innovation bias shifts towards

the labor factor. Quantitatively, the innovation response is large as the shift in the innovation bias ranges between 19.4 p.p. (for transport equipment) and -11.5 p.p. (for textiles).

In a next step, we assess the export response to the REE supply shock. Panel (b) of Figure 6 depicts the change in industry-level exports for countries other than China. For a better comparison, we normalize exports with the global exports of the respective industry. While industry exports fall throughout all industries in response to the REE supply restrictions, it is noteworthy that this decline is minimal throughout all industries. More importantly and consistent with the empirical estimates, the more REE-intensive industries contract relatively less.

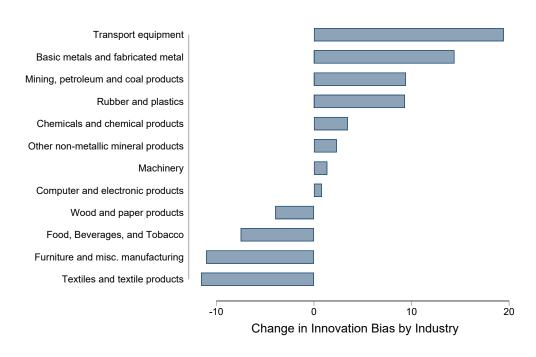
Impact on GDP and Welfare: We now turn to a discussion of the general-equilibrium effects of the Chinese REE policies. In Figure 7, we show the impact on real GDP. The global response of value added highlights the importance of directed technological change in cushioning the response to the Chinese REE policies. While there have been real GDP losses outside of China, these losses were very small throughout, in the range of 0.002% to 0.004%. The Chinese gain in real GDP due to the policy is also fairly small, around 0.02%.

We use our model quantification to infer the welfare consequences of China's policy. We measure welfare in terms of real consumption of the final good. Thus, changes in real value added need to be adjusted for primary incomes from the tradable REE factor input and for REE tax revenues to obtain real gross national expenditures of each region on the non-tradable final good. We deduct investment for innovation from gross national expenditures to obtain real consumption. Consequently, while more innovation raises GDP, it also reduces the part of gross national expenditures that is available for consumption due to the presence of innovation fix costs. The REE policies increased China's real consumption by about 0.2% while they had negligible welfare costs on regions outside China. In the U.S., the consumption decline was around 0.006% and the consumption decline in Europe and Japan was even smaller.

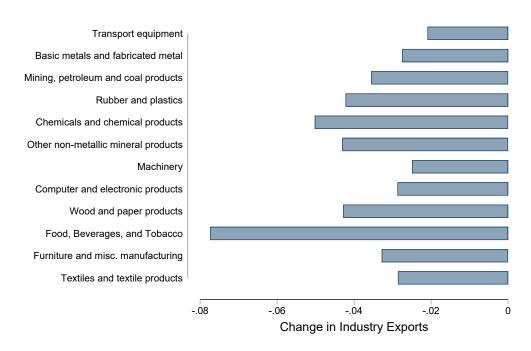
Fixed Technologies To highlight the role of directed technological change to cushion the general-equilibrium impact of the REE supply shock, we contrast these results with a counterfactual scenario under fixed technologies. Under this counterfactual, we examine the impact that the same Chinese policies would have had on real GDP if technologies had remained fixed at their initial levels, without allowing for technological change. In such a counterfactual world, as Figure 7 shows, real GDP losses outside of China would have been sizable. The real GDP loss in the U.S. would have been above 0.5%, in Europe and Japan around 0.3%. Conversely, China would gain much more from the policy in a world without endogenous technology response than under directed technological change. Holding technologies fixed, the restrictive REE policies would have caused a gain of more than 1.8% in real GDP for China.

Figure 6: Directed Technological Change and Exports

#### (a) Directed Technological Change

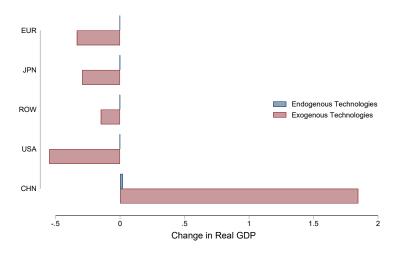


#### (b) Exports



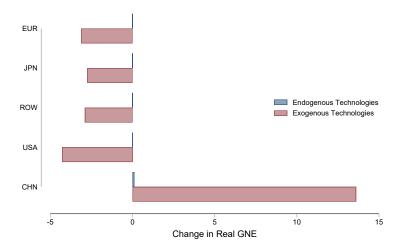
Notes: The figure plots relative changes in  $A_{Ris}/A_{Lis}$  outside of China (measured in % change from the baseline economy) in response to the introduction of the Chinese REE export tax. Regions are weighted according to real GDP in the baseline economy.

Figure 7: Effect on Real GDP



Notes: The figure plots relative changes in real GDP (measured in % change from the baseline economy) in response to the introduction of the Chinese REE policies.

Figure 8: Effect on Gross National Expenditures



Notes: The figure plots relative changes in real gross national expenditures (measured in % change from the baseline economy) in response to the introduction of the Chinese REE policies.

With exogenously fixed technologies, these relative losses in real consumption would have been much larger, with around 5.0% in the U.S. and more than 3.0% in Europe and Japan (Figure 8). Since a direct comparison of welfare effects between the factual model and the counterfactual under fixed technologies would be biased due to a lack of investments under fixed technologies, as there are no innovation fixed costs, we compare differences in real gross national expenditures in Figure 8. Changes in gross national expenditures would also have been a lot larger under fixed technologies, with China gaining more than 10% and other regions losing between 2.8% and 4.3%.

Increased REE Usage Finally, we use the structure of our model to study the counterfactual GDP impact of the identical Chinese REE policies in a world that uses more REEs. By 2025, the production and usage of REEs has approximately doubled compared to levels before the REE crisis. Hence, we consider the effect of the 40% supply restriction in combination with an export tax rate of  $\tau_{XC} = 8$  in a world with a twice as large REE endowment. While the effects on real GDP under endogenous technologies are similarly small, the impact of that policy under exogenous technologies would have been higher. It would have caused a GDP gain of 2.1% for China (compared to 1.8%) and a loss of 0.6% for the U.S. (compared to 0.5%) and losses between 0.3% and 0.4% for Europe and Japan (compared to 0.3%).

# 7 Conclusion

How did export restrictions on REEs affect downstream industries both globally and domestically? To address this question, we construct a detailed IO table that accounts for individual REE inputs, allowing us to assess their significance for specific industries. Our empirical analysis provides robust evidence that the export restrictions imposed by China between 2010 and 2015 stimulated a strong innovation response in REE-intensive downstream industries outside of China. To measure REE-biased innovation, we classify patents from downstream industries across the world based on whether they substitute for or enhance the efficiency of REE usage. Our findings further indicate a significant increase in total factor productivity and exports in the most exposed industries outside of China, relative to less affected industries. Furthermore, we do not observe a comparable increase in total factor productivity or exports within the respective Chinese industries.

To quantify the general-equilibrium effects of China's export restrictions, we then develop a novel quantitative trade model that incorporates directed technological change. We calibrate the model using trade and production data. Leveraging the model structure, we estimate the elasticity of substitution between REEs and labor, finding that these inputs are complementary in most industries. Consistent with our reduced-form evidence, our model predicts that an increase in the international REE price - caused by a Chinese export tax and supply restriction - induces

technological change aimed at reducing REE usage, that is sufficiently strong to offset the impact of rising input costs. In contrast to a counterfactual economy with exogenously fixed technologies, we show that the endogenous innovation response outside of China has prevented large real GDP and welfare losses while the positive GDP and welfare impact on China was largely dampened compared to a counterfactual world with fixed technologies.

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# A Empirical Appendix

# A.1 Imputing REEs into the Supply-Use Table

We first convert REE use numbers from the USGS report (Bleiwas and Gambogi, 2013), which are in metric tons, to USD million using a combination of prices from BCC (2015) and Asian Metal, both at the element level. BCC reports global consumption of REEs by element, both in metric tons and USD million units, giving us the prices per unit. For elements not reported by BCC, we impute their prices by extrapolating from Asian Metal, which reports prices in Chinese markets. We also extrapolate the numbers to 2012 using the compounded annual growth rate of overall rare earth consumption from 2010 to 2012, which is based on data from USGS Mineral Commodity Summaries.

For imputation into the supply-use table, we match each "general category" of REE content from the USGS report into its corresponding NAICS code, which the BEA's SUT uses. Table A.2 presents this matching, which determines which using sectors (columns) we assign the REE use numbers to. For instance, we match the "Fluid Cracking Catalysts" category with "Petroleum Refineries" column. Unfortunately, not all categories of REEs applications can be neatly matched to NAICS. This is an issue especially for the category of permanent magnets, which is the largest in metric tons and USD value. The closest match is the NAICS code for "Other fabricated metal manufacturing." While this NAICS code includes permanent magnets, it also includes industrial pattern manufacturing, steel wool manufacturing, and various others.

To alleviate this lack of granularity in the input sector "Other fabricated metal manufacturing", we split it into magnet and non-magnet production. We then draw from the list of the final use of permanent magnets from Alonso et al. (2023) and BCC (2015) to assign NAICS codes, which would use inputs from the newly split magnet sector, while the rest would take inputs from the artificial non-magnet sector. Table A.3 presents the list of magnet-using sectors along with the split weights. For example, the NAICS sector "Turbine and turbine generator set units manufacturing" takes USD 55 million of inputs from "Other fabricated metal manufacturing". Since we assigned this sector as magnet-using with full weight in our IO table, it would take all USD 55 million of inputs from "Other fabricated metal manufacturing – Magnets" and none from "Other fabricated metal manufacturing – Non-magnets".

Meanwhile, we assign REEs as an input to the NAICS code for "Other Basic Inorganic Chemical Manufacturing" for the supplying sectors (rows). This is the closest match for rare earth oxides, which are the form of REE inputs reported in the USGS report. We split this NAICS code into six rows: five for the individual REEs specified in our raw data source and one for non-REEs. We then impute the numbers into the corresponding supply-use pairs, e.g. the USD value of Nd for magnets

is assigned to the cell with the supplying NAICS of "Other Basic Inorganic Chemical Manufacturing – Nd" and the using NAICS of "Other fabricated metal manufacturing – Magnets". We leave the diagonal cells for the split REE compound sectors empty, e.g., "Other Basic Inorganic Chemical Manufacturing – Nd" does not use inputs supplied by itself or from the other REE compounds. As for the column and row totals, we split them using proportions from the values of each REE oxides approximated from USGS.

Using the imputed supply-use table, we compute total requirement of the supplying REE sector, carved out from "Other Basic Inorganic Chemical Manufacturing", by each using NAICS industry. As a last step, we convert these numbers to the SIC-level by using concordance mapping from Pierce and Schott (2012) as the outcomes we study vary at the SIC (and country) level.

# A.2 Index of Complementarity

We use the substitute performance index developed by Graedel et al. (2015) to account for whether REE inputs used by the sector are highly complementary in the production function. The concept of complementarity here arises from exogenous physical and chemical properties of the elements. The index is constructed by listing potential substitutes for each element's primary uses and then assessing their performance as informed by the assimilation of research and expert opinion.

For example, for Cerium (CE), the authors analyze different applications (e.g., "Glass polishing"), application details (e.g., "Used to polish precision optics"), percentage of application (e.g., 25% global use), primary substitute (e.g., iron oxide), and substitute performance (e.g., "adequate"). For instance, in the application of glass polishing—which accounts for 25% of global cerium use—the primary substitute is iron oxide, with an "adequate" performance rating. However, in battery alloys (10%), cerium is replaced by lithium-ion batteries, rated as a "good" substitute, reflecting greater ease of replacement. For "other" uses like arc welding and carbon arc lighting (16%), no substitute is identified, and the performance is marked as "not applicable." By assigning numerical scores to performance ratings—such as "adequate" or "good"—and weighting them by the percentage of cerium use in each application, the author quantifies overall substitutability. This results in a composite index that reflects the difficulty of replacing cerium across its various industrial roles.

There is considerable variation among REE elements, with Dysprosium being the least substitutable and samarium being the most (Figure A.2).

#### A.3 Patent Data

Our primary measure for shifts in the direction of technological change is based on patent data. We obtain the universe of granted patents related to REE from the Google Patent Research database.

REE Patents: As a first step, we identify patents as broadly related to REE and link them to individual elements when their title or abstract contains certain keywords that include either the name of the elements themselves, their chemical compounds, or some key related technologies, such as technologies related to permanent magnets.<sup>22</sup> By the end of the sample period, there exist around 30,000 granted unique REE patents, globally. We assign the country of the patent based on the country of the patent assignee, considering the same set of top 50 countries as in the other analyses that we pool to the regions Europe, U.S. and Canada, China, Russia, Korea, Japan, Australia and the Rest of World.

Classifying Patents in a LLM: In the next step, we link each patent in our sample to an industry. This is not a straightforward task, as patents are categorized in technology classes instead of industries. To assign a patent to an industry, researchers have previously created and used concordance tables linking technology classes to their sectors of use.<sup>23</sup>

While this method is in principle feasible for our classification, a concern is that REE-related technologies are generally much younger than concordances between IPC technology classes and the SIC industry classification. Furthermore, technologies usually map into large groups of industries, and this mapping is not specific to REE. Alternatively, patents have been linked to industries via the industry affiliation of the firm holding the patent.<sup>24</sup> Linking patents via firms has two caveats. First, a large fraction of REE patents is filed by non-corporate entities such as universities and other research institutions or by non-public firms. Neglecting these patents could bias our results if REE-related research were systematically more prevalent in the corporate or non-corporate sectors. Second, while patent-firm matches exist for some countries (such as the NBER/Compustat patent matches for the U.S.), a global matched patent-firm database is not readily available.

Instead of following these paths, we use an LLM to assign the patents in our sample to industries. For that purpose, we parse the title and abstract text of each individual patent to the LLM and let the LLM suggest the SIC industry that fits best to the patent. For this approach, we use the model GPT4 from OpenAI. We refine the sample of REE patents by asking the LLM to classify

<sup>&</sup>lt;sup>22</sup>See Table A.5 for the list of keywords.

<sup>&</sup>lt;sup>23</sup>See, e.g., Silverman (2002).

 $<sup>^{24}</sup>$ See e.g. Bloom et al. (2013).

whether each patent describes a technology that improves the use of REEs or helps to substitute away from the use of REE.

Other Patents: We draw a large random sample of patents and let the LLM allocate the SIC manufacturing industry codes. We then scale these numbers for each year across regions by the observable number of granted patents that each region has in the total sample of manufacturing patents for that year.

## A.4 Constructing a Country-Industry Productivity Panel

The primary data sources to construct a comprehensive panel dataset for TFP and labor productivity across countries and industries are the United Nations Industrial Development Organization (UNIDO) INDSTAT and OECD STAN databases. UNIDO INDSTAT provides detailed industry-level data on economic indicators such as value added, employment, and capital formation, available at both the 3-digit and 4-digit ISIC levels for Revisions 3 and 4. ISIC Rev. 3 offers extensive historical coverage up to 2008, while Rev. 4 provides improved coverage from 2008 onwards. All monetary values are expressed in current US dollars. Due to the unbalanced nature of the 4-digit INDSTAT data, substantial imputation is necessary to construct a balanced panel. The process involves nearest-neighbor interpolation to fill missing observations within the 4-digit data wherever possible. Annual growth rate series at the more consistently available 2-digit ISIC level are computed and used to adjust the imputed data at the 4-digit level Remaining gaps are filled using 2-digit level data that are split into 4-digit industries based on time-invariant employment shares. This imputation process is performed separately for both versions of INDSTAT (ISIC Rev. 3 and Rev. 4) to maintain consistency with their respective coverage periods.

We derive initial capital stock estimates at the country-industry level from OECD STAN. When initial capital stock data are missing, estimates are obtained through nearest-neighbor interpolation or – if still missing – a regression approach using gross capital formation as a predictor. For the following sample years, we use the perpetual inventory method and data on gross capital formation to obtain a measure of the capital stock, assuming a depreciation rate of 8%. We deflate capital stocks and other variables in nominal U.S. dollars using the price deflators for capital and consumer prices from the Penn World Tables vers. 9.1.

TFP is calculated as the residual from an OLS regression estimated at the 2-digit ISIC level, where the log of value added is regressed against the logs of capital stocks and the log of employment. Labor productivity is derived as the log difference of value added and employment.

Finally, we map the ISIC productivity data to SIC manufacturing industries using 2 concordances

mapping 4-digit ISIC Rev. 3 or ISIC Rev. 4 to SIC codes. The mapping from ISIC Rev. 4 is done via ISIC Rev. 3.1 as an intermediate step. In cases where mappings are not unique, ISIC industries are weighted based on time-invariant employment weights. Only productivity data at the U.S. SIC level where less than one-third of the data is imputed is used to ensure data integrity and reliability.

For U.S. and Japan, we use data from their respective manufacturing survey or census in place of the UNIDO INDTSTAT database. For the former, we use the NBER CES manufacturing database, which have shipment, inputs and productivity variables as well as price deflators readily available at the 4-digit SIC level. We construct a similar dataset for Japan by assembling the country's annual Manufacturing Census data from the Ministry of Economy, Trade and Industry (METI) from the years 2002 to 2018. METI's data comprises of output (shipment and production) and input (number of employees, salaries expense, raw materials usage, and tangible fixed assets) variables, as in the NBER CES manufacturing database, at the 4-digit JSIC level for all establishments with 30 or more employees. This covers 521 manufacturing industries after combining the eleventh, twelfth and thirteenth JSIC revisions.

To obtain real values for Japan's productivity measures, we make use of price indices from the Bank of Japan's Input-Output Price Index (IOPI) dataset. For raw material prices, we use input weights from Japan's 2005 input-output table to calculate the weighted average of input price indices for each input-output sector classification. Since some inputs from the input-output table have no matching price indices from IOPI, we impute with Japan's producer price index for total intermediate goods. Meanwhile for service inputs, we use price indices from BoJ's Japan Services Producer Price Index dataset. We build a concordance between Japan's input-output sector classification with JSIC to merge the price indices into the manufacturing census data.

We then calculate TFP measures as residual of the factor inputs employees, capital, and raw materials, imitating the approach employed by the NBER CES manufacturing database for comparability. Specifically, we subtract from output growth the growth in the number of employees, raw materials excluding energy, energy, and fixed assets, each weighted by their share in output in Yen value. Finally, we map the JSIC-level data to SIC manufacturing industries using our manually-constructed concordance.

### A.5 Trade and Other Data

We use data on exports of 4-digit SIC level manufacturing industries for the 50 largest economies in the world from 2002 until 2018 from UN Comtrade.

We construct a country-industry proxy of how strongly an industry is targeted by industrial policy

using the subsidy database from the Global Trade Alert (GTA). This proxy is the share of subsidies (counted in the number of policy measures) going to a country-industry relative to the total number of subsidy policies worldwide.

## A.6 Results with Alternative REE Sensitivity Measures

For our regressions of exports growth, we consider an alternative construction of REE sensitivity where industry-level total requirements of REE, now aggregated across elements, are multiplied with country-level initial share of REE imports from China instead of with element-level complementarity index. We use UN Comtrade to construct alternative measures of exposure to the Chinese REE supply shock that vary at the country-industry level-considering HS codes 284690, 284610, and 280530— and weighing them by initial import shares, using:

1. Country-level price spike in unit values of REE imports interacted with the total requirements of REE of each industry (aggregated across elements):

$$REE \ Sensitivity_{is} = \left(\sum_{e} tr_{es}\right) \times \left(\ln(\max \ REE \ import \ price_i) - \ln(REE \ import \ price \ 2009_i)\right).$$

2. Pre-shock import shares of REEs in country c sourced from China relative to REEs sourced from a larger set of countries (China, the U.S., Australia, Russia or India) interacted with the total requirement of REEs in industry s:

$$REE \ Sensitivity_{is} = \left(\sum_{e} tr_{es}\right) \times \frac{REE \ imports \ from \ CHN_i}{REE \ imports \ from \ CHN, \ USA, \ AUS, \ RUS, \ IND_i).}$$

Table A.8 shows that export growth of downstream industries increased relatively more in those country-industries where the REE supply shock was relatively more important.

#### A.7 Results at HS Product Code Level

We also run our downstream export growth regressions at the HS product code level instead of SIC as in the baseline. The purpose of this is to assess the robustness of our findings as well as to decompose the effect on exports into its quantity and unit price components. The latter is possible since the raw UN Comtrade data, before converting to SIC-level, also include the physical quantity of exports. We calculate unit prices by simply dividing exports value with exports quantity. We

exclude HS codes for which there was a change in the unit of physical quantity, but this applies to less than 5% of observations for our sample period. For the REE sensitivity variable, we reconstruct it at the HS-level by converting NAICS-level total requirements to HS using the Pierce-Schott concordance instead of the SIC concordance. We make use of the same concordance to convert our control variables.

For exports value, we find that the results are qualitatively similar to our baseline regressions at the SIC level. Specifically, for all our sub-samples except for U.S. and China, exports of product codes that are more sensitive to REE inputs exhibit significantly larger growth during the treatment period compared to before than exports of less-sensitive products by the same country. The same regressions for exports quantity and unit value confirm that the effects for non-China countries are driven more by quantity of exports than just by passing through higher prices. This is demonstrated by the coefficients for non-China and all countries in the exports unit value regressions being not significant or less so than in the exports quantity regressions. For European countries and Japan, where the positive exports effect is most pronounced, the coefficient magnitudes are similarly larger for physical quantity.

# A.8 Examples of Patents

This appendix presents selected REE-related patents from our dataset to exemplify the innovation response to the REE supply-shock episode. These are patents filed after 2010 whose content could have the effect of reducing or substituting the use of REEs as inputs. They have also been handpicked to represent the variety of REE downstream uses.

Table A.1: List of Sample Countries in Empirical Analysis

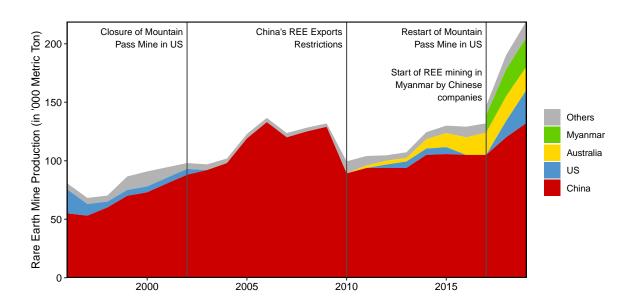
ISO Code	Country Name	ISO Code	Country Name
ARE	United Arab Emirates	ISR	Israel
ARG	Argentina	ITA	Italy
AUS	Australia	$_{ m JPN}$	Japan
AUT	Austria	KOR	South Korea
$\operatorname{BEL}$	Belgium	MEX	Mexico
BGD	Bangladesh	MYS	Malaysia
BRA	Brazil	NGA	Nigeria
CAN	Canada	NLD	Netherlands
CHE	Switzerland	NOR	Norway
$\operatorname{CHL}$	Chile	NZL	New Zealand
$\operatorname{CHN}$	China	PAK	Pakistan
$\operatorname{COL}$	Colombia	PER	Peru
CZE	Czech Republic	PHL	Philippines
DEU	Germany	POL	Poland
DNK	Denmark	PRT	Portugal
EGY	Egypt	ROM	Romania
ESP	Spain	RUS	Russia
FIN	Finland	SAU	Saudi Arabia
FRA	France	$\operatorname{SGP}$	Singapore
GBR	United Kingdom	SWE	Sweden
HKG	Hong Kong	THA	Thailand
IDN	Indonesia	TUR	Turkey
IND	India	USA	United States of America
$\operatorname{IRL}$	Ireland	VNM	Vietnam
IRN	Iran	ZAF	South Africa

Notes: The table lists the sample countries in our empirical analysis, which are the 50 largest economies based on their GDP.

Table A.2: Categories of Rare-Earth Applications from the U.S. Geological Survey

General Category (USGS)	NAICS Code	NAICS Description
Alloys Batteries Automobile catalyst Fluid catalytic cracking Magnets	331110 335911 336390 324110 332999	Iron and steel mills and ferroalloy manufacturing Storage battery manufacturing Other motor vehicle parts manufacturing Petroleum refineries Other fabricated metal manufacturing (split into magnet and non-magnet)

Figure A.1: Rare-Earth Mine Production Across Countries



Notes: The figure plots production of REEs for major producing economics using data from the Mineral Yearbooks of the U.S. Geological Survey.

Table A.3: NAICS Codes Designated as Magnet-Using

NAICS Code	NAICS Description	Weight
333415	Air conditioning, refrigeration, and warm air heating equipment manufacturing	1.0000
333611	Turbine and turbine generator set units manufacturing	1.0000
333613	Mechanical power transmission equipment manufacturing	1.0000
334112	Computer storage device manufacturing	1.0000
334118	Computer terminals and other computer peripheral equipment manufacturing	1.0000
334510	Electromedical and electrotherapeutic apparatus manufacturing	1.0000
334610	Manufacturing and reproducing magnetic and optical media	1.0000
335222	Household refrigerator and home freezer manufacturing	1.0000
335312	Motor and generator manufacturing	1.0000
335314	Relay and industrial control manufacturing	1.0000
336111	Automobile manufacturing	1.0000
336310	Motor vehicle gasoline engine and engine parts manufacturing	1.0000
336320	Motor vehicle electrical and electronic equipment manufacturing	1.0000
336350	Motor vehicle transmission and power train parts manufacturing	1.0000
336390	Other Motor Vehicle Parts Manufacturing	1.0000
336411	Aircraft manufacturing	1.0000
336414	Guided missile and space vehicle manufacturing	1.0000
339112	Surgical and medical instrument manufacturing	1.0000
339114	Dental equipment and supplies manufacturing	1.0000
339910	Jewelry and silverware manufacturing	1.0000
33391A	Pump and pumping equipment manufacturing	1.0000
33441A	Other electronic component manufacturing	1.0000
3363A0	Motor vehicle steering, suspension component (except spring), and brake systems manufacturing	1.0000
33641A	Propulsion units and parts for space vehicles and guided missiles	1.0000
332913	Plumbing fixture fitting and trim manufacturing	0.2500
333111	Farm machinery and equipment manufacturing	0.2500
333120	Construction machinery manufacturing	0.2500
333517	Machine tool manufacturing	0.2500
333618	Other engine equipment manufacturing	0.2500
334513	Industrial process variable instruments manufacturing	0.2500
336413	Other aircraft parts and auxiliary equipment manufacturing	0.2500
33299A	Ammunition, arms, ordnance, and accessories manufacturing	0.2500
33329A	Other industrial machinery manufacturing	0.2500
33399A	Other general purpose machinery manufacturing	0.2500
334111	Electronic computer manufacturing	0.0625
334300	Audio and video equipment manufacturing	0.0625
336412	Aircraft engine and engine parts manufacturing	0.0625

Notes: The table lists the weights we use for splitting the NAICS code for "Other fabricated metal manufacturing" into "magnets" and "non-magnets". For each NAICS code that uses USD x million input from "Other fabricated metal manufacturing" and has weight w, in the imputed input-output table it would take USD xw million from "magnets" and USD x(1-w) million from "non-magnets".

**Table A.4:** Rare Earth Total Requirements ( $10^{-3}$  USD of REE per 1 USD of SIC Final Demand)

No SIC	Description	All	Ce	La	Nd	Pr	Dy
1 3691	Storage Batteries					0.00	
2 3499	Fabricated Metal Products, NEC	5.91	0.00	0.00	4.06	0.18	1.66
$3\ 3625$	Relays and Industrial Controls	0.58	0.00	0.00	0.40	0.02	0.16
$4\ 3511$	Turbines and Turbine Generator Sets	0.53	0.00	0.00	0.36	0.02	0.15
5 3292	Asbestos Products	0.47	0.01	0.00	0.32	0.01	0.13
6 3714	Motor Vehicle Parts and Accessories	0.41	0.09	0.00	0.22	0.01	0.09
7 3519	Internal Combustion Engines, NEC	0.39	0.19	0.00	0.14	0.01	0.06
8 3585	Refrigeration and Heating Equipment	0.37	0.18	0.00	0.13	0.01	0.05

Figure A.2: Index of Complementarity by Element

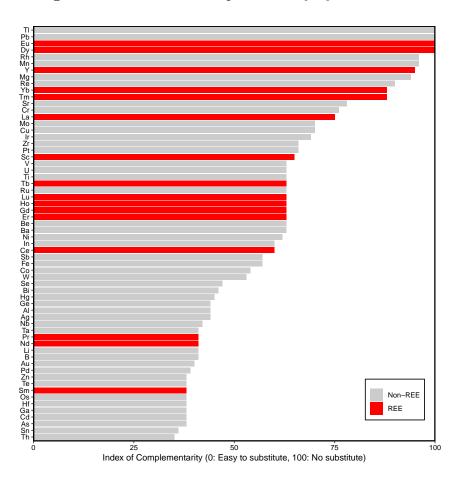


Table A.5: Keywords for the Rare-Earth Patent Search

Element	Keywords	Element	Keywords
Cerium	cerium ceo2	Praseodymium	prnd ndfeb
Dysprosium	dysprosium dy2o3		rare earth magnet rare-earth magnet
Erbium	erbium er2o3		rare earth element magnet rare-earth element magnet
Gadolinium	gadolinium gd2o3		nib magnet neo magnet
Holmium	holmium		nd2fe14b
	ho2o3	Scandium	scandium
Lanthanum	lanthanum		sc2o3
	la2o3	Samarium	samarium
Lutetium	lutetium lu2o3		sm2o3 smco
Neodymium	neodymium nd2o3 ndfeb rare earth magnet		rare earth magnet rare-earth magnet rare earth element magnet rare-earth element magnet
	rare-earth magnet rare earth element magnet	Terbium	terbium tb4o7
	rare-earth element magnet nib magnet	Yttrium	yttrium y2o3
	neo magnet nd2fe14b	Ytterbium	ytterbium yb2o3
	prnd	Europium	europium
Praseodymium	praseodymium		eu2o3
	pr2o3		

*Notes:* The table lists keywords used to search for REE patents from Google Patent Research database. Eu, Pm, and Tm were excluded due to being too rare for most industrial applications.

**Table A.6:** Patents in Rare-Earth Intense Manufacturing Industries, Robustness using Country-Level Rare-Earth Import Intensity

		REE-E	Enhancing F	Patents	
	NONCHN	ALL	USA	EUR	JPN
	(1)	(2)	(3)	(4)	(5)
	REE	Sens.is base	d on REE i	mport price	surge
REE Sens. $\times$ Post	426.1***	425.4***	363.9***	624.8***	921.5**
	(100.6)	(102.4)	(124.0)	(185.2)	(373.8)
Observations	5,561	7,606	1,200	1,140	972
Clusters	387	531	81	74	66
		REE-E	Enhancing F	Patents	
	NONCHN	ALL	USA	EUR	JPN
	(6)	(7)	(8)	(9)	(10)
	REE Sens	.is based on	the REE in	nport share f	from CHN
REE Sens. $\times$ Post	1143.0***	1075.3***	1201.6**	1222.8***	1849.2***
	(320.8)	(318.4)	(471.8)	(409.7)	(691.5)
Observations	5,561	7,606	1,200	1,140	972
Clusters	387	531	81	74	66
Controls	Yes	Yes	Yes	Yes	Yes
Country $\times$ Ind F.E.	Yes	Yes	Yes	Yes	Yes
Country $\times$ Year F.E.	Yes	Yes	Yes	Yes	Yes

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{rst} = \beta REE Sens._s \times post_t + \gamma \Delta_{rst} + \eta_{rs} + \eta_{rt} + \epsilon_{rst}$ with Poisson pseudo-maximum likelihood estimation. The outcome  $y_{rst}$  represents granted REE-related patents that improve the efficiency of REE or help find ways to substitute REE usage. The sample includes 4-digit SIC manufacturing industries (with at least one REE-related patent) from 2002-2018 across 8 regions. Regions capture the location of the patent assignee and include Australia, China, European Union, Korea, Russia, Japan, U.S. and the Rest of the World. In the upper panel, REE Sens. is measures the impact of the REE import price surge in country i and is calculated as the product of the total REE requirement share in industry s and the REE import price spike in country i. The price spike is defined as the logarithmic difference between the peak weighted REE import price (across REE-related HS codes 280530, 284690 and 284610, typically occurring between 2011 and 2013) and the average REE import price in 2016.  $REE\ Sens._{is} = (\sum_{e} tr_{es}) \times (\ln(max\ REE\ import\ price_i) - \ln(REE\ import\ price\ 2009_i)).$ In the bottom panel, REE Sens. is based on the average share of REE imports from China before the Chinese export restrictions and is calculated as the product of the total REE requirement share in industry s and the proportion of REE imports from China relative to total REE imports from China, the USA, Australia, Russia, and India in i between 1995 and 2009. REE Sens. is  $(\sum_e tr_{es}) \times (REE \ imports \ from \ CHN_i/REE \ imports \ from \ CHN, \ USA, \ AUS, \ RUS, \ IND_i).$  post<sub>t</sub> is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, and Japan. All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes timeinvariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and country-industry-specific industrial subsidy fractions from the Global Trade Alert database, all interacted with  $post_t$  and the lagged weighted average growth rate of GDP of the ten largest importers from is. Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

**Table A.7:** Productivity Growth of Rare-Earth Intense Manufacturing Industries, Robustness using Country-Level Rare-Earth Import Intensity

	Annual	lized Grow	th: Total	Factor Pr	oductivity			
	NONCHN (1)	ALL (2)	USA (3)	EUR (4)	JPN (5)			
	REE S	$Sens{is}$ bas	ed on RE	E import p	rice surge			
REE Sens. $\times$ Post	16.17* (8.672)	15.07* (9.030)	12.92* (7.127)	18.48** (8.230)	37.76*** (12.63)			
Observations Clusters	177,577 14,272	$182,791 \\ 14,646$	$5{,}320$ $380$	$88,055 \\ 6,306$	$\frac{4,978}{363}$			
	Annual	Annualized Growth: Total Factor Productivity						
	NONCHN (6)	ALL (7)	USA (8)	EUR (9)	JPN (10)			
	REE Sens.i	s based on	the REE	import sh	are from CHN			
REE Sens. $\times$ Post	40.05* (22.02)	41.46* (22.49)	38.25* (20.52)	46.36** (19.60)	80.53*** (27.85)			
Observations Clusters	183,818 14,981	$\substack{189,032 \\ 15,355}$	5,320 380	88,055 6,306	4,978 363			
	Anr	nualized G	rowth: La	abor Produ	ıctivity			
	NONCHN (11)	ALL (12)	USA (13)	EUR (14)	JPN (15)			
	REE S	Sens.is bas	ed on RE	E import p	rice surge			
REE Sens. $\times$ Post	9.945 (6.161)	9.372 (6.344)	3.391 (13.94)	2.489 (5.113)	38.67*** (14.90)			
Observations Clusters	177,577 14,272	182,791 14,646	5,320 380	88,055 6,306	5,685 383			
	Anr	nualized G	rowth: La	abor Produ	ictivity			
	NONCHN (16)	ALL (17)	USA (18)	EUR (19)	JPN (20)			
	REE Sens.i	s based on	the REE	import sh	are from CHN			
REE Sens. $\times$ Post	31.12** (13.96)	32.23** (14.16)	10.40 (39.64)	17.11 (12.74)	86.25*** (32.32)			
Observations Clusters	183,818 14,981	$189,032 \\ 15,355$	$5,320 \\ 380$	88,055 6,306	$\frac{5,685}{383}$			
Country × Ind F.E. Country × Year F.E.	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes	Yes Yes Yes			

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{ist} = \beta \ REE \ Sens._s \times post_t + \gamma \Delta_{ist} + \eta_{is} + \eta_{it} + \epsilon_{ist}$ , where  $y_{ist}$ represents the annualized growth rate of productivity for country-industry is in year t, TFP (upper two panels) or labor productivity measured as value added per worker (lower two panels). The annualized growth is calculated using the midpoint between t and t-1as the denominator. The sample includes 4-digit SIC manufacturing industries from 2002-2018 across the 50 largest economies. In the first and third panels, REE Sens. is measures the impact of the REE import price surge in country i and is calculated as the product of the total REE requirement share in industry s and the REE import price spike in country i. The price spike is defined as the logarithmic difference between the peak weighted REE import price (across REE-related HS codes 280530, 284690 and 284610, typically occurring between 2011 and 2013) and the average REE import price in 2016. REE Sens.  $is = (\sum_e tres) \times (\ln(max\ REE\ import\ price_i) - (\ln(max\ REE\ import\ price_i))$  $\ln(REE\ import\ price\ 2009_i)$ ). In the second and fourth panels,  $REE\ Sens._{is}$  is based on the average share of REE imports from China before the Chinese export restrictions and is calculated as the product of the total REE requirement share in industry s and the proportion of REE imports from China relative to total REE imports from China, the USA, Australia, Russia, and India in i between 1995 and 2009. REE Sens.  $i_s = (\sum_e tr_{es}) \times (REE \text{ imports from } CHN_i/REE \text{ imports from } CHN, USA, AUS, RUS, IND_i).$  post<sub>t</sub> is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, and Japan. For the country subsamples of the U.S. and Japan, we use the data from NBER-CES Manufacturing Database and Japan's Annual Manufacturing Census, respectively (see Appendix A.4 for notes). All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and country-industry-specific industrial subsidy fractions from the Global Trade Alert database, all interacted with post, and the lagged weighted average growth rate of GDP of the ten largest importers from is. Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

**Table A.8:** Downstream Export Growth of Rare-Earth Intense Manufacturing Industries, Robustness using Country-Level Rare-Earth Import Intensity

	A	nnualized	Growth: 1	Exports Valu	ue		
	NONCHN	ALL	USA	EUR	JPN		
	(1)	(2)	(3)	(4)	(5)		
	REE S	Sens. <sub>is</sub> based	d on REE	import pric	ce surge		
REE Sens. $\times$ Post	25.92***	26.38***	7.516	29.88***	52.50*		
	(8.088)	(8.020)	(18.64)	(11.18)	(26.74)		
Observations	166,701	172,684	6,048	78,044	5,979		
Clusters	$10,\!500$	10,874	378	4,884	375		
	A	Annualized Growth: Exports Value					
	NONCHN	ALL	USA	EUR	JPN		
	(6)	(7)	(8)	(9)	(10)		
	REE Sens.	s based on	the REE	import share	e from CHN		
REE Sens. $\times$ Post	75.99***	77.50***	20.35	78.18***	119.1**		
	(20.59)	(20.59)	(53.98)	(26.79)	(53.74)		
Observations	178,330	184,313	6,048	78,044	5,979		
Clusters	11,236	11,610	378	4,884	375		
Controls	Yes	Yes	Yes	Yes	Yes		
Country $\times$ Ind F.E.	Yes	Yes	Yes	Yes	Yes		
Country $\times$ Year F.E.	Yes	Yes	Yes	Yes	Yes		

Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{ist} = \beta REE Sens._{is} \times post_i + \gamma \Delta_{ist} + \eta_{is} + \eta_{it} + \epsilon_{ist}$ , where  $y_{ist}$ represents the annualized growth rate of export values for country-industry is in year t. The annualized growth is calculated using the midpoint between t and t-1 as the denominator. The sample includes 4-digit SIC manufacturing industries from 2002-2018 across the 50 largest economies, excluding China. In the upper panel, REE Sens.is measures the impact of the REE import price surge in country i and is calculated as the product of the total REE requirement share in industry s and the REE import price spike in country i. The price spike is defined as the logarithmic difference between the peak weighted REE import price (across REE-related HS codes 280530, 284690 and 284610, typically occurring between 2011 and 2013) and the average REE import price in 2016. REE Sens.  $is = (\sum_e tr_{es}) \times tr_{es}$  $(\ln(\max{REE~import~price_i}) - \ln(REE~import~price~2009_i))$ . In the bottom panel,  $REE~Sens._{is}$  is based on the average share of REE imports from China before the Chinese export restrictions and is calculated as the product of the total REE requirement share in industry s and the proportion of REE imports from China relative to total REE imports from China, the USA, Australia, Russia, and India in i between 1995 and 2009.  $REE\ Sens._{is} = (\sum_e tr_{es}) \times (REE\ imports\ from\ CHN_i/REE\ imports\ from\ CHN,\ USA,\ AUS,\ RUS,\ IND_i).$ post<sub>t</sub> is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, and Japan. All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, and country-industry-specific industrial subsidy fractions from the Global Trade Alert database, all interacted with  $post_t$  and the lagged weighted average growth rate of GDP of the ten largest importers from is. Regressions are weighted by the share of export values  $w_{is} = exports_{is}/\sum_{i.s} exports_{is}$  from 2002-2004. Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01.

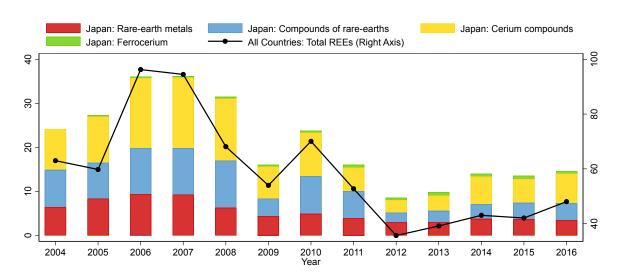
**Table A.9:** Downstream Export Growth of Rare-Earth Intense Manufacturing Industries, HS-Level Data

	Annualized Growth: Exports Value							
	NONCHN	ALL	USA	EUR	JPN	CHN		
	(1)	(2)	(3)	(4)	(5)	(6)		
REE Sens. $\times$ Post	1.532***	1.533***	0.543	1.682***	1.617*	1.434		
	(0.312)	(0.303)	(1.040)	(0.412)	(0.938)	(1.397)		
		Annualized Growth: Exports Quantity						
	NONCHN	ALL	USA	EUR	JPN	CHN		
	(7)	(8)	(9)	(10)	(11)	(12)		
REE Sens. $\times$ Post	1.164***	1.083***	-0.753	1.129***	1.536	-0.103		
	(0.347)	(0.335)	(0.984)	(0.403)	(1.330)	(1.123)		
		Annualized	Growth: E	xports Unit	Value			
	NONCHN	ALL	USA	EUR	JPN	CHN		
	(13)	(14)	(15)	(16)	(17)	(18)		
REE Sens. $\times$ Post	0.390	0.473*	1.918***	0.481	0.0987	1.584		
	(0.246)	(0.242)	(0.520)	(0.302)	(0.770)	(1.194)		
Observations	2,016,181	2,075,321	53,871	865,220	53,423	59,140		
Clusters	$157,\!752$	$161,\!850$	4,003	$64,\!376$	4,028	4,098		
Controls	Yes	Yes	Yes	Yes	Yes	Yes		
Country $\times$ Ind F.E.	Yes	Yes	Yes	Yes	Yes	Yes		
Country $\times$ Year F.E.	Yes	Yes	Yes	Yes	Yes	Yes		

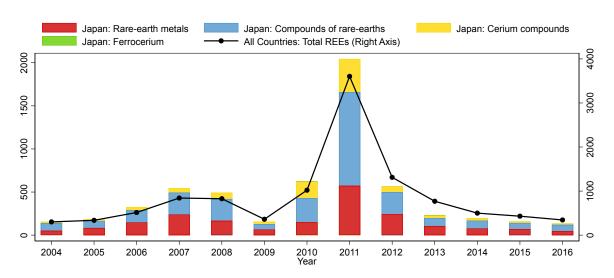
Notes: The table presents coefficient estimates of  $\beta$  from the regression:  $y_{ist} = \beta REE \ Sens._s \times post_t + \gamma \Delta_{st} + \eta_{is} + \eta_{it} + \epsilon_{ist}$ , where  $y_{ist}$  represents the annualized growth rate of export values, quantity and unit price for country-industry is in year t. The annualized growth is calculated using the midpoint between t and t-1 as the denominator. The sample includes 6-digit HS product codes from 2002-2018 across the 50 largest economies. The treatment intensity  $REE \ Sens._s$  is a weighted sum of an REE element-specific complementarity index (ranging from 0 to 100), with weights based on total requirement shares for industry s:  $REE \ Sens._s = \sum_e tr_{es} \times compl_e$ .  $post_t$  is a dummy variable set to 1 for 2010 and later years (post-China's REE export restrictions). Country subsamples include non-China, all countries, the U.S., European economies, Japan, and China. All regressions include country-industry and country-year fixed effects. The control vector  $\Delta_{ist}$  includes time-invariant measures of industry s's capital and labor intensity from the NBER CES manufacturing database, all interacted with  $post_t$ . Standard errors (in parentheses) are clustered at the country-industry level. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01.

Figure A.3: Annual REE Imports From China

#### (a) Quantity (in metric ton)



## (b) Value (in USD million)



*Notes:* The figure presents import values of REEs from China by Japan and all countries, broken down by subcategories of REE commodities for the former. Data is sourced from UN Comtrade.

Figure A.4: US8480815B2: Magnet Powder Coating Process

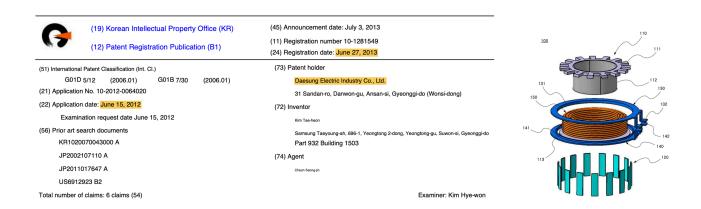
#### (12) United States Patent US 8,480,815 B2 (10) Patent No.: DETAILED DESCRIPTION OF THE INVENTION (45) Date of Patent: Jul. 9, 2013 Wang Magnets made using the present process use much less Dy or Tb than those made using the conventional methods while METHOD OF MAKING ND-FE-B SINTERED OTHER PUBLICATIONS obtaining similar magnetic properties. In the present process, MAGNETS WITH DY OR TB Plusa, et al., Domain Structure and Domain-Wall Energy in Polycrystalline R2Fe14B Compounds (R=Pr, Nd, Gd, Dy), Journal the Dy or Tb coated Nd—Fe—B powders are used to make the magnet, which results in a non-uniform distribution of Dy (75)Inventor: Yucong Wang, West Bloomfield, MI of the Less-Common Metals, 1987, pp. 231-243, vol. 133, The Nethor Tb in the magnet, which can be seen and measured using a erlands. Rodewald, Magnetization and Aging of Sintered Nd-Fe-B Magnets, Journal of the Less-Common Metals, 1985, pp. 77-81, vol. 111, The Netherlands. Machida, et al., Improved Magnetic Properties of Small-Sized NdFe-B Magnets and Their Application for DC brush-less Micro-Motors, Center for Advanced Science and Innovation, Osaka University, scanning electron microscope with a microprobe. This GM Global Technology Operations (73) Assignee: enables the present process to use much less Dy or Tb for the similar magnetic properties. For example, the amount of Dy Subject to any disclaimer, the term of this (\*) Notice: and/or Tb can be reduced by about 20% or more compared to patent is extended or adjusted under 35 U.S.C. 154(b) by 58 days. conventional processes, or about 30% or more, or about 40% May 2, 2005, pp. 25-30, Japan. Herget, et al., Metallurgical Methods for the Production of Rare Earth-Transition Metal Permanent Magnet Materials, MPR Jun. 1987, pp. 438-444. or more, or about 50% or more, or about 60% or more, or (21) Appl. No.: 13/007,203 about 70% or more, or about 80% or more, or about 90% or more. By non-uniform distribution, we mean that Dy and/or Japanese Office Action with a due date of Apr. 21, 2013 pertaining to Jan. 14, 2011 (22) Filed: JP Application No. 2011-245644 The are distributed or concentrated at the interface of the HIGH VOLUME FLOW GRANULATION LIQUID LOW VOLUME FLOW Assuming the weight of permanent magnet pieces is about 5 1-1.5 kg per electric motor, and a yield of the machined permanent magnet (PM) pieces of typically about 55-60%,

Assuming the weight of permanent magnet pieces is about 5 1-1.5 kg per electric motor, and a yield of the machined permanent magnet (PM) pieces of typically about 55-60%, 2-3 kg of PM per motor would be required, or 4-6 kg per vehicle (some hybrid vehicles may use one induction motor and one PM motor). Moreover, Dy is also widely used by 5 other industries. The only RE mine in the United States does not have any significant amounts of Dy. Therefore, reducing the Dy or Tb usage in permanent magnets would have a very significant cost impact.

240 225 225 225 226 220 220 220 220 220

Notes: The figure shows excerpts of patent US8480815B2. In January 2011, the US firm GM Global Technology Operations filed a patent for a powder coating process that uses much less Dy and Tb as coating for REE magnets. The patent claims to reduce the amount of Dy and Tb used by at least 20% while maintaining similar magnetic properties. The patent cites supply constraints as motivation as "the only RE mine in the United States does not have any significant amounts of Dy", which could be an implicit reference to the difficulties of importing these elements from China.

Figure A.5: KR101281549B1: Position Sensors Without REE



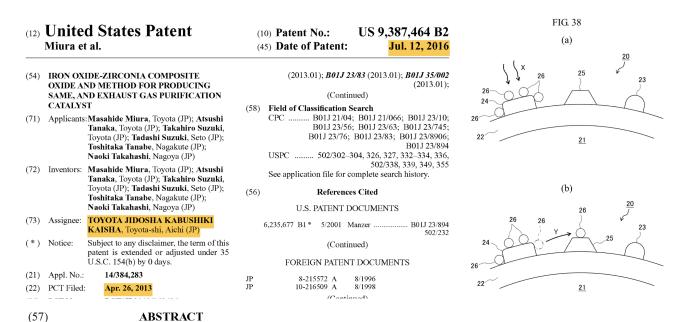
The collected magnetic flux may be transmitted to the magnetic sensor (not shown) through the house terminals 33 and 34 to detect the twist between the axes by the manipulation of the handle.

However, in the conventional position sensor structure, although the sensor output can be sufficiently generated by the detection of the desired magnetic flux amount, there is a cost problem due to the recent increase of the rare earth price because the permanent magnet of rare earth material is used.

Accordingly, it is an object of the present invention to provide a position sensor having an improved structure so that a permanent magnet of rare earth material is not used.

Notes: The figure shows excerpts of patent KR101281549B1. Around one year after REE prices peaked, the Korean firm Daesung Electric Co filed a patent in South Korea for position sensors with a modified structure that removes the need for permanent REE magnets. The patent cites the positive price shock as its motivation, stating that "there is a cost problem due to the recent increase of the rare earth price". Position sensors have many downstream applications, from manufacturing processes to transport equipments including automobiles.

Figure A.6: US9387464B2: Catalyst for Exhaust Gas Purification



A composite oxide with a high oxygen storage capacity is provided without using cerium. The composite oxide is an iron oxide-zirconia composite oxide containing iron, zirconium, and a rare-earth element. The total content of Fe<sub>2</sub>O<sub>3</sub>, ZrO<sub>2</sub>, and an oxide of the rare-earth element is not less than 90 mass %, the content of an iron oxide in terms of Fe<sub>2</sub>O<sub>3</sub> is 10 to 90 mass %, and the absolute value of the covariance COV (Fe, Zr+X) of the composite oxide, which has been baked in the atmosphere at a temperature of greater than or equal to 900° C. for 5 hours or more, is not greater than 20.

ite oxide obtained by causing an iron oxide to be supported on a support containing ceria.

Cerium contained in such composite oxides is expensive, and a problem has emerged that cerium is now difficult to obtain stably due to the deterioration of the procurement environment in recent years. Thus, suppressing the amount of cerium used is considered.

However, it is recognized by one of ordinary skill in the art that when the content of cerium is reduced in a composite

Notes: The figure shows excerpts of patent US9387464B2. In April 2016, Toyota Motor from Japan filed a patent for a composite oxide used in catalysts for exhaust gas purification. This type of catalysts uses mainly Cerium, the REE with the greatest price jump in 2010-2011. The patented composite oxide still needs to use REEs but it does not have to be Cerium. In the example, it preferably uses Lanthanum and Yttrium. Similar to the other patent examples, the patent mentions recent supply issues in its background: "Cerium contained in such composite oxides is expensive, and a problem has emerged that Cerium is now difficult to obtain stably due to the deterioration of the procurement environment in recent years." This patent comes only a few years after the REE supply shock but builds on recent patents from before the supply shock: "Meanwhile, JP 2008-93496 A (Patent Literature 5) discloses a promoter clathrate containing an iron oxide, which is a promoter of an exhaust gas purification catalyst, and a zirconia solid solution (e.g., Example 2). In such a promoter clathrate, the iron oxide is covered with the zirconia solid solution. Thus, sintering of the iron oxide is suppressed, and consequently, an exhaust gas purification catalyst containing such a promoter clathrate exhibits excellent catalyst activity."

# B Model Appendix

# B.1 Derivation of an Estimation Equation for $\varepsilon_s$

Combining (15) and relative demand (41) yields

$$\frac{A_{Ris}}{A_{Lis}} = \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\varepsilon_s} \left(\frac{P_{Ris}}{P_{Lis}}\right)^{1 - \varepsilon_s} \frac{f_{Li}}{f_{Ri}} = \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\varepsilon_s} \left(\frac{A_{Ris}}{A_{Lis}}\right)^{(1 - \varepsilon_s)(\frac{1}{1 - \mu_s} - \delta)} \left(\frac{w_{Ri}}{w_{Li}}\right)^{1 - \varepsilon_s} \frac{f_{Li}}{f_{Ri}}, \quad (A.1)$$

where the last equality uses  $P_{Ris} = A_{Ris}^{\frac{1}{1-\mu_s}-\delta} \frac{\mu_s}{1-\mu_s} w_{Ri}$  and the corresponding expression for  $P_{Lis}$ . Solving for  $A_{Ris}/A_{Lis}$ , we obtain an expression of relative patents as a function of relative factor prices  $w_{Ri}/w_{Li}$ :

$$\frac{A_{Ris}}{A_{Lis}} = \left(\frac{\gamma_s}{1 - \gamma_s}\right)^{\frac{\varepsilon_s(\mu - 1)}{\kappa_s}} \left(\frac{w_{Ri}}{w_{Li}}\right)^{\frac{(1 - \varepsilon_s)(\mu_s - 1)}{\kappa_s}} \left(\frac{f_{Li}}{f_{Ri}}\right)^{\frac{\mu_s - 1}{\kappa_s}}.$$
(A.2)

Taking logs, we obtain our regression specification:

$$\log\left(\frac{A_{Ris}}{A_{Lis}}\right) = \beta_s \log\left(\frac{w_{Ri}}{w_{Li}}\right) + \delta_s + u_{is}. \tag{A.3}$$

# B.2 Algorithm to Calibrate the Model to a Baseline Economy

We calibrate the model to an initial baseline economy, using the following fixed-point iteration routine. An outer loop iterates over fixed costs  $f_{Ris}$  for given  $f_{Lis}$  until the model matches  $\frac{A_{Ris}}{A_{Lis}}$  to relative patents in the data for 2009, according to equation (15). Within that outer loop, the algorithm iterates over an inner loop that adjusts  $Rev_{is}$  to be consistent with clearing product markets according to equation (25), iterates over trade imbalances to match country-level GDP for the given  $Rev_{is}$  and calculates the according factor expenditures. Once relative patents and GDP are matched, the algorithm moves on to prices and trade costs. First, it solves for the set of bilateral prices  $P_{ijs}$  between each exporters and importers in each sector that exactly reproduces the observed trade shares  $\lambda_{ijs}$  and calculates the according price indices  $P_{is}$  and  $P_i$ . With these prices and the model-implied unit costs of production, it then infers iceberg trade costs  $d_{ijs}$ . Lastly, it calibrates the demand weights  $\alpha_{is}$  so that, when combined with the sectoral prices, they replicate the observed spending patterns and the overall price level in each country. At this point, the baseline economy is fully specified.

# B.3 Algorithm to Solve the Model

For given parameter values, the model can be solved by the following fixed-point iteration routine. An outer loop iterates over  $Rev_{is}$  to clear product markets, according to equation (25). Within that loop, first innovation responses and factor shares are calculated, for given levels of  $Rev_{is}$ . Second, we can directly solve for factor prices to clear factor markets as described in equations (35) and (36). Third, we iterate over the system of bilateral prices  $P_{ijs}$  to be consistent with (29), (30) and (31). Fourth, we calculate transfers  $T_i$  according to (27).